

Phase Transitions in Semidefinite Relaxations (a fast & robust algorithm for community detection)

Federico Ricci-Tersenghi
(Sapienza University)

in collaboration with
Adel Javanmard and Andrea Montanari

arxiv:1511.08769 → PNAS 2016
arXiv:1603.09045 → J. Phys. Conf. Ser. 2016

Outline of the talk

- Hidden partition model (a.k.a. community detection pb.)
 - Spectral vs. optimization methods (ML & SDP)
 - Optimality vs. robustness
- Community detection algorithm based on SDP
 - It is simple, fast and robust
- Phase transitions in SDP
 - Statistical physics approach -> phase transitions
 - (Quasi-)optimality of SDP for (sparse) dense graphs

Communities detection problem

- Detecting communities/partitions/clusters in graphs is a widespread problem in many different disciplines
- Examples of applications: social networks mining, recommendation systems improvement, images segmentation and classification, and many more in biology...
- We need fast (linear and scalable) algorithms
 - robust (real datasets are very noisy and not random)
 - optimal (on random ensemble benchmarks)

Benchmark for community detection

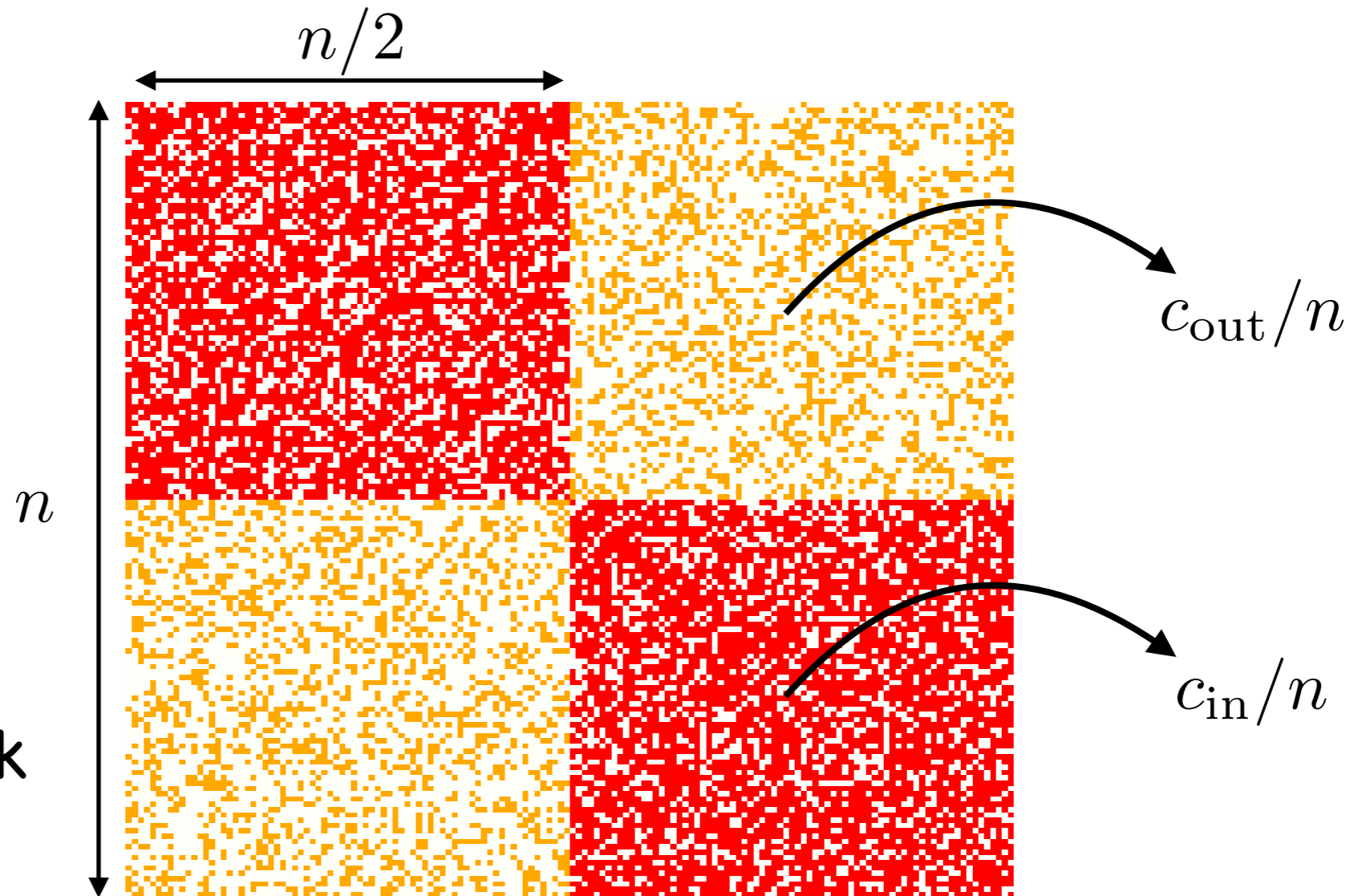
Hidden partition model or stochastic block model (SBM)

- Generate a partition of n nodes: e.g. q groups of size n/q
- Add independently edges between any pair of nodes according to the following probability

$$\mathbb{P}[(ij) \in E] = \begin{cases} c_{\text{in}}/n & \text{same group} \\ c_{\text{out}}/n & \text{different groups} \end{cases}$$

- Assortative model $c_{\text{in}} > c_{\text{out}}$
Disassortative model $c_{\text{in}} < c_{\text{out}}$

The hidden partition model



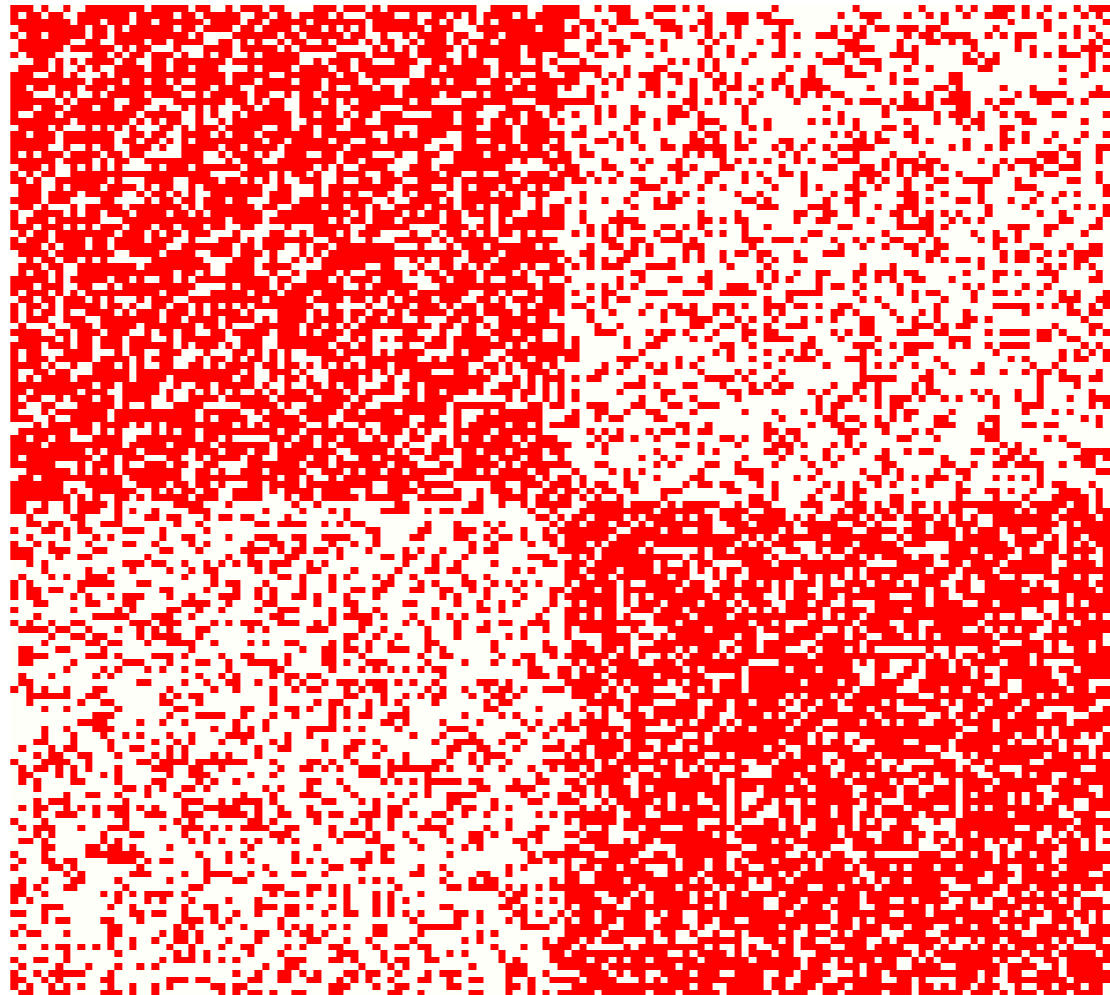
Stochastic block
model (SBM)
with $q = 2$

$$\mathbb{P}[(ij) \in E] = \begin{cases} c_{in}/n & \text{same group} \\ c_{out}/n & \text{different groups} \end{cases}$$

Assortative model:

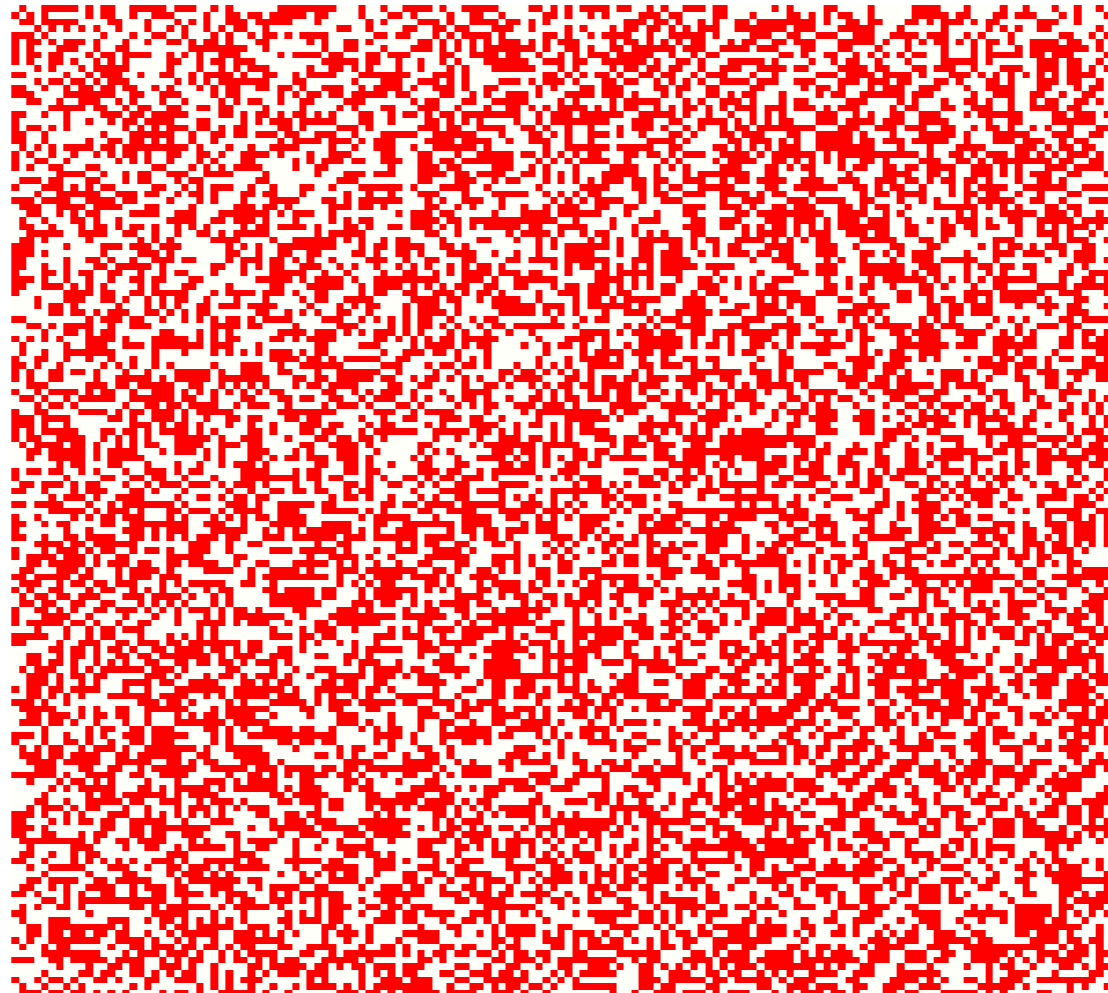
$$c_{in} > c_{out}$$

The hidden partition model



Colors are not provided !

The hidden partition model

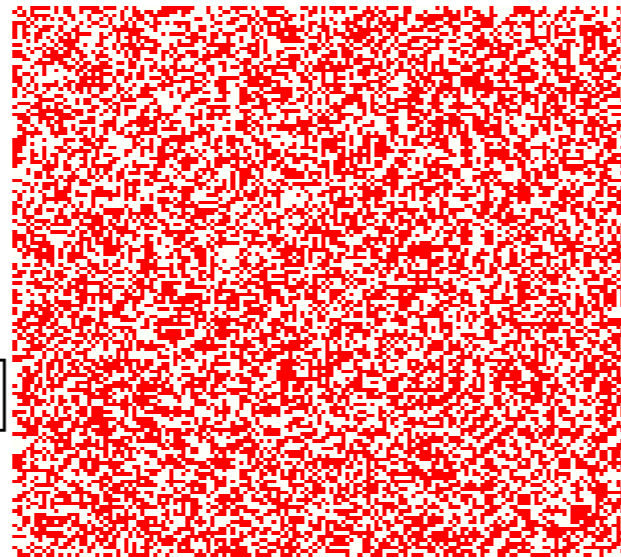


The right ordering neither !!

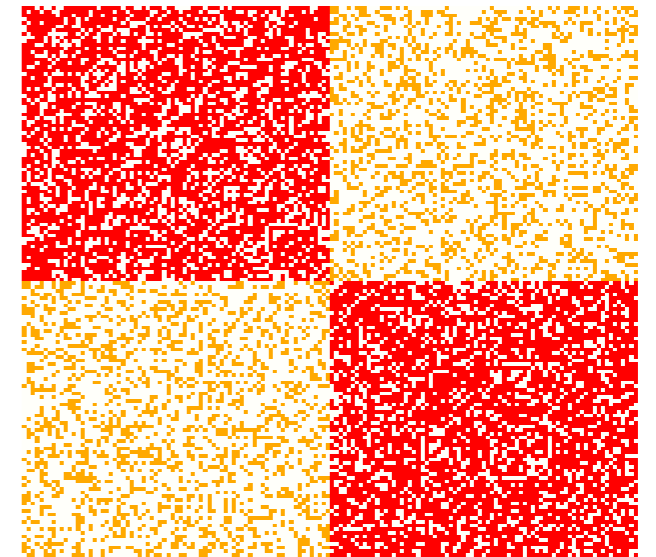
The hidden partition model

Given only the adjacency matrix

$$A_{ij} = A_{ji} = \mathbb{I}[(ij) \in E]$$



Infer the
hidden
partition



Hidden (true) partition $\rightarrow \mathbf{x}_0 \in \{+1, -1\}^n$

Estimated partition $\rightarrow \hat{\mathbf{x}}(G) \in \{+1, -1\}^n$

Quality of inference

via the overlap $\rightarrow Q = \frac{1}{n} |\langle \hat{\mathbf{x}}(G), \mathbf{x}_0 \rangle|$

Assortative SBM with 2 equal-size groups

Relevant parameters and threshold

- Mean degree $d = \frac{c_{in} + c_{out}}{2}$
- Signal-to-noise ratio $\lambda = \frac{c_{in} - c_{out}}{2\sqrt{d}}$

- Bayes optimal threshold $\lambda_c = 1$
- Impossible detection for $\lambda < \lambda_c$
- BP algorithm with $Q > 0$ for $\lambda > \lambda_c$

Very
ingenious
spectral
methods

[Decelle, Krzakala, Moore, Zdeborova, 2011]
[Massoulié, 2013] [Mossel, Neeman, Sly, 2013]

Maximum Likelihood (ML)

- If no information on the generative model is given (apart being assortative and with 2 equal-size groups) a good choice is to maximize the likelihood

$$\text{maximize } \sum_{i,j} A_{i,j} x_i x_j$$

$$\text{subject to } x_i \in \{+1, -1\} \text{ and } \sum_i x_i = 0$$

- NP-hard problem

Lagrangian formulation

- Maximize $\sum_{i,j} A_{ij} x_i x_j - \eta \left(\sum_i x_i \right)^2$ over $x \in \{+1, -1\}^n$
- A good choice is $\eta \geq d/n$
- For $\eta = d/n$ the centered adjacency matrix appears

$$A_{ij}^{\text{cen}} = A_{ij} - d/n$$

- Maximize $\sum_{i,j} A_{ij}^{\text{cen}} x_i x_j$ over $x \in \{+1, -1\}^n$

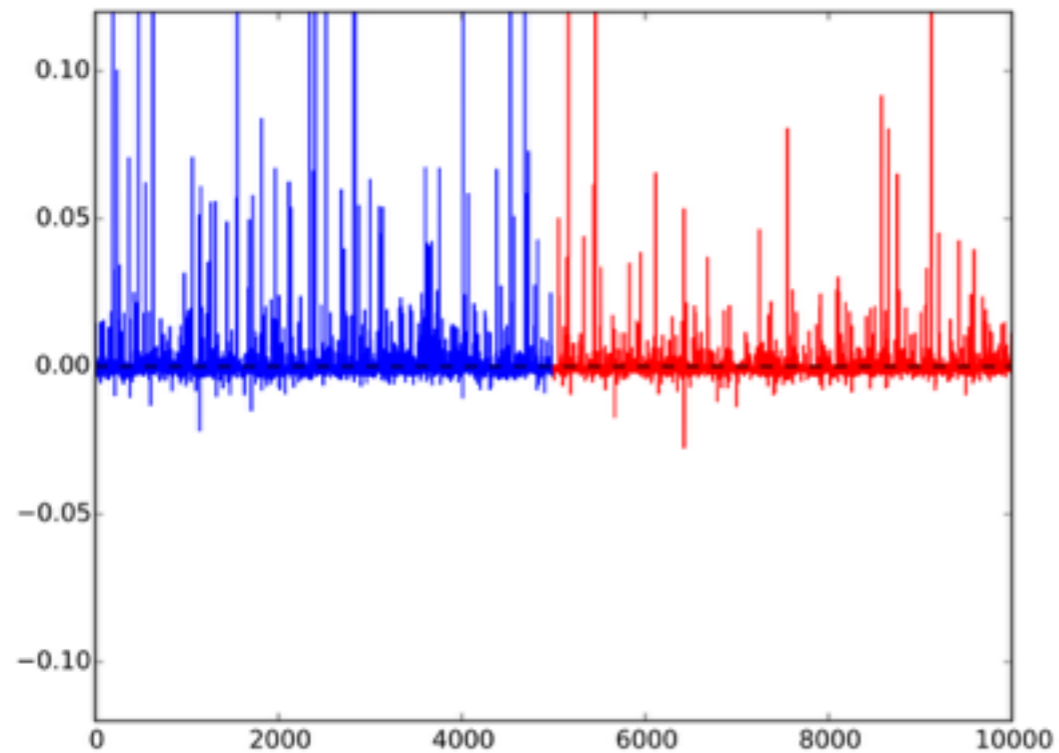
Spectral relaxation (PCA)

- PCA on A^{cen} relaxes the constraint $x \in \{+1, -1\}^n$ and maximizes $\sum_{i,j} A_{ij}^{\text{cen}} x_i x_j$ over $x \in \mathbb{R}^n$
- Computes the eigenvector of the largest eigenvalue $v_1(A^{\text{cen}})$
- Estimates the partition via a projection on $x \in \{+1, -1\}^n$
$$\hat{x}^{\text{PCA}} = \text{sign}(v_1(A^{\text{cen}}))$$
- It is good as long as components of $v_1(A^{\text{cen}})$ have similar moduli/intensities

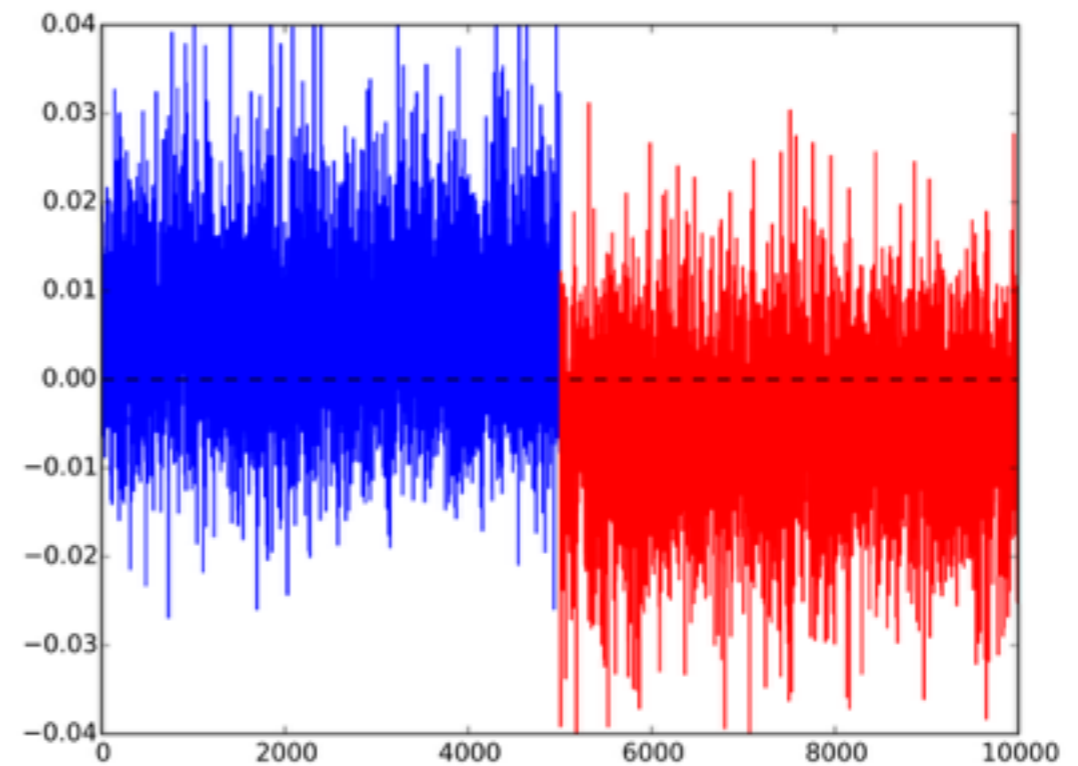
Spectral relaxation fails on sparse graphs

$$n = 10^4 \quad \lambda = 1.2$$

$$d = 3$$



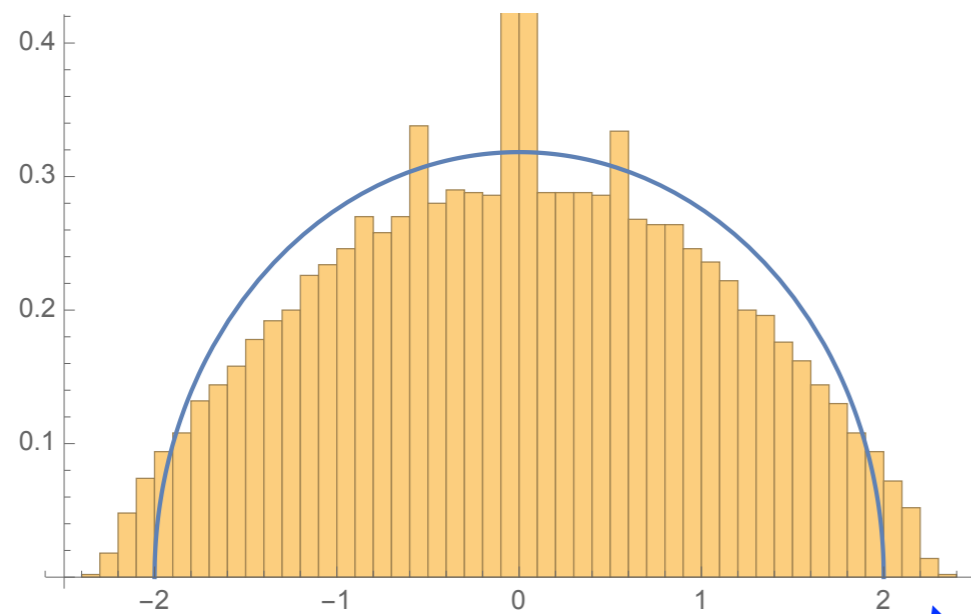
$$d = 20$$



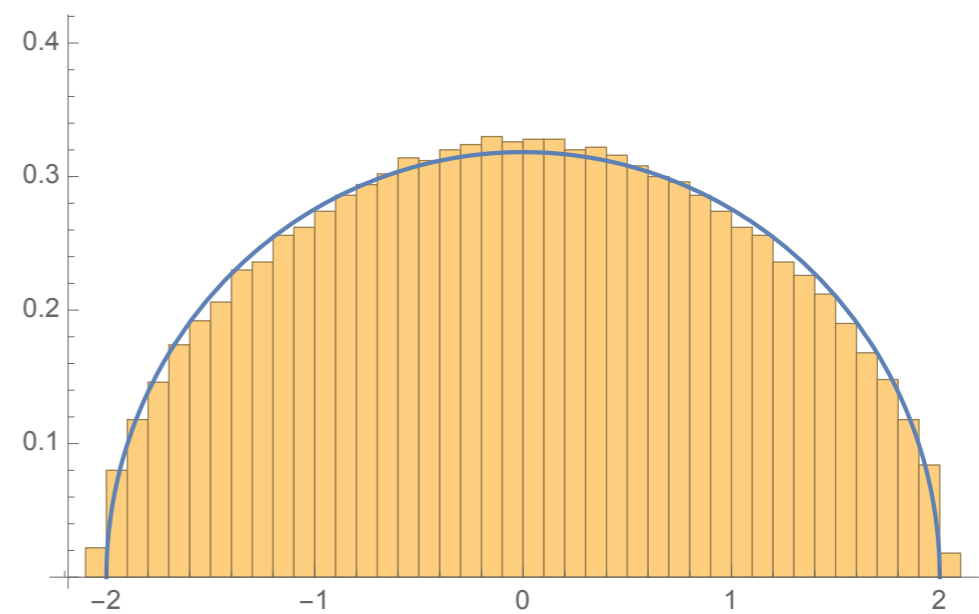
$$v_1(A^{\text{cen}})$$

Why PCA fails?

$$\frac{1}{\sqrt{d}} A^{\text{cen}} = \frac{\lambda}{n} x_0 x_0^T + W$$



$d = 3$



$d = 20$

localized
eigenvectors

Other spectral methods

- Compute eigenvalues and eigenvectors of some matrix related to the adjacency matrix A
- Fail for the same reason -> eigenvectors localization
- Laplacian $L = D - A$
with D being the diagonal matrix of degrees
- Normalized Laplacian $\mathcal{L} = D^{-1/2} L D^{-1/2}$
- Shown to be sub-optimal in the sparse regime
[Kawamoto, Kabashima, 2015]

$$\lambda_c^{\text{Lap}} = \sqrt{\frac{d}{d-1}} > 1$$

Improved spectral methods

- Non-backtracking matrix

[Krzakala, Moore, Mossel, Neeman, Sly, Zdeborova, Zhang, 2013]

- seems to avoid localization around large degree nodes
- optimal for the SBM
- complex spectrum, not easy to compute

- Bethe Hessian [Saade, Krzakala, Zdeborova, 2014]

$$H(r) = (r^2 - 1)\mathbb{1} - rA - D$$

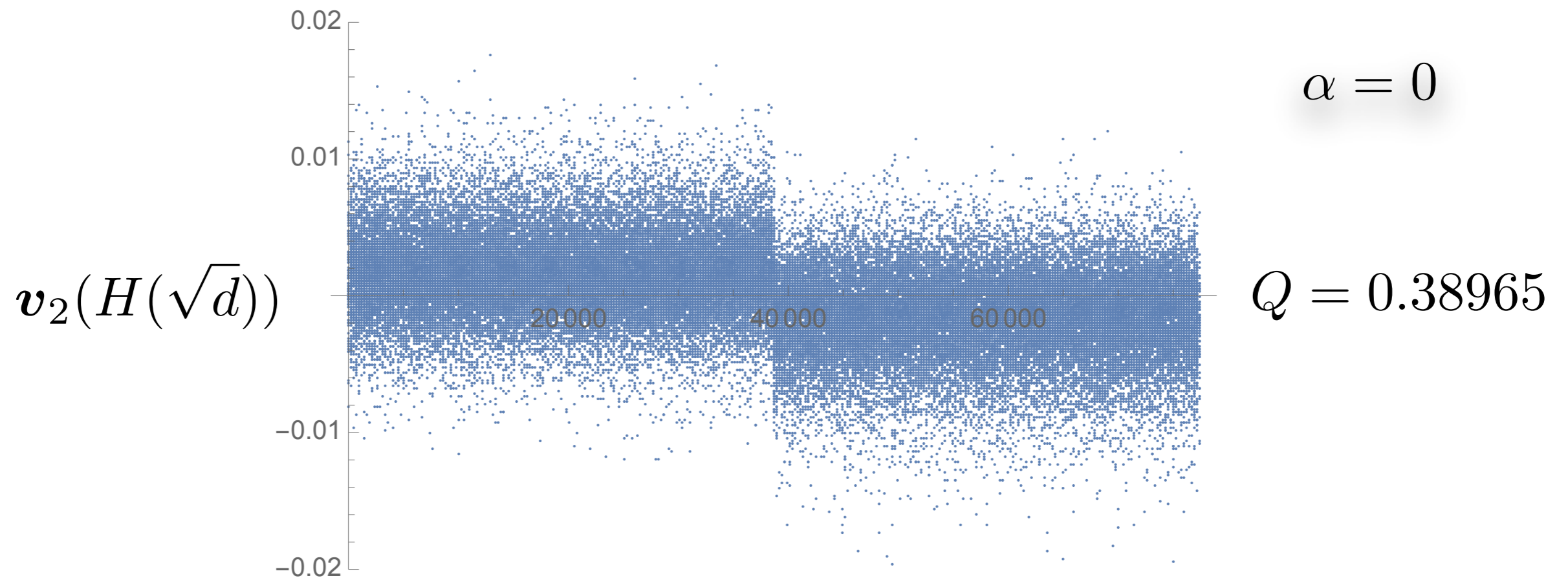
- symmetric matrix, real spectrum
- easier to use
- optimal for the SBM with $r = \sqrt{d}$

Quasi-random graphs

- Generate a graph according to the SBM
- Choose a subset S of vertices of size $|S| = \alpha n$
- For each vertex in S connect all its neighbours
- The number of edges increases by $\sim \alpha d^2 n / 2$
i.e. by a fraction $\sim \alpha d$
- A robust inference method should work also for $\alpha > 0$
at least in the regime $\alpha \ll 1/d$

Improved spectral methods fail on quasi-random graphs

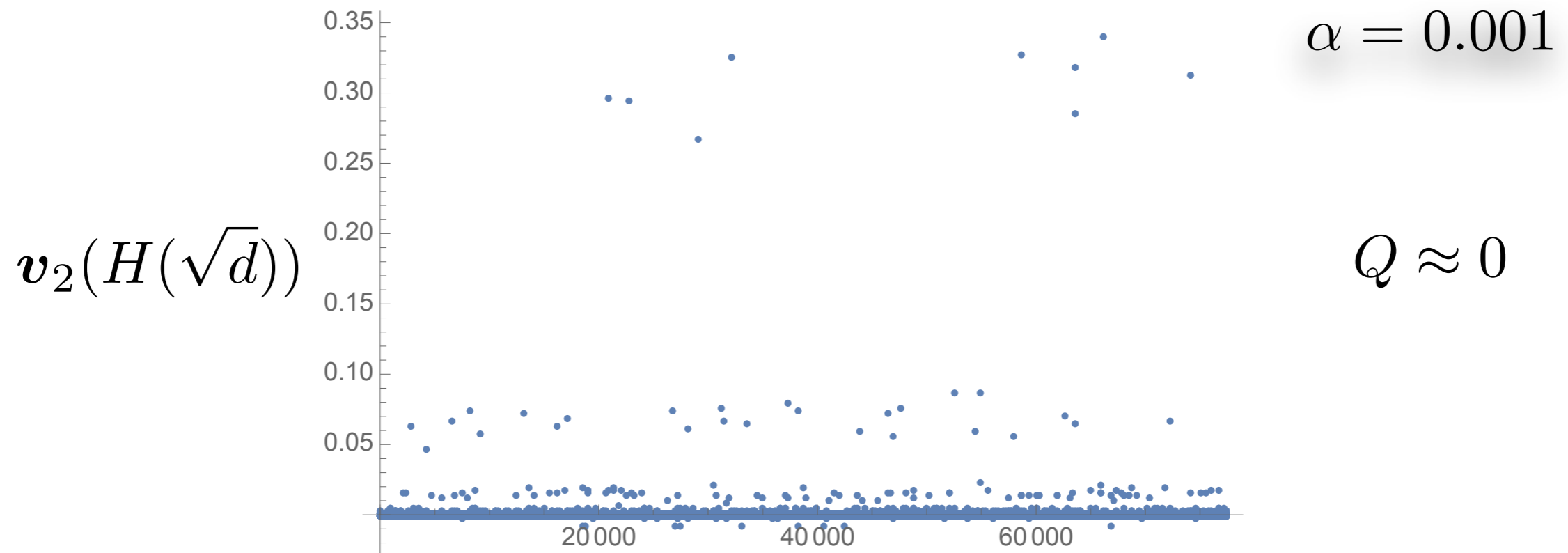
$$n = 10^5 \quad d = 3 \quad \lambda = 1.1$$



(on the 2-core 77336 nodes and 132763 edges)

Improved spectral methods fail on quasi-random graphs

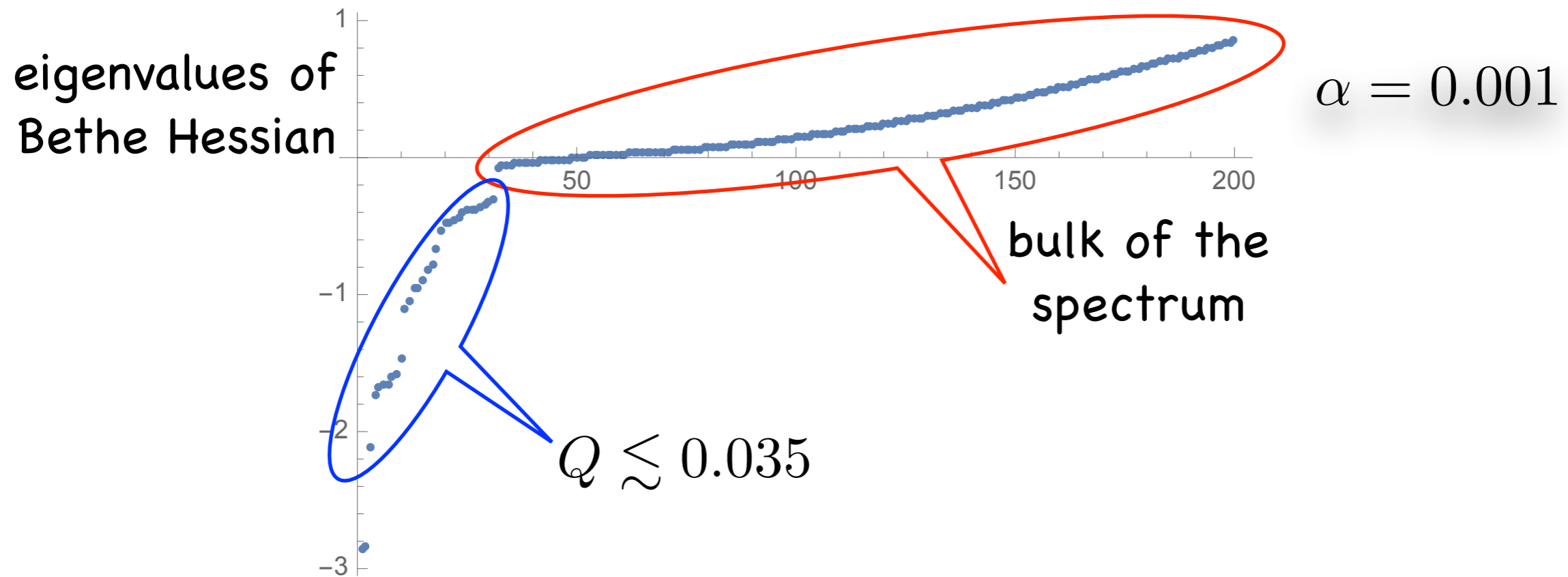
$$n = 10^5 \quad d = 3 \quad \lambda = 1.1$$



(on the 2-core 77336 nodes and 133185 edges)

Improved spectral methods fail on quasi-random graphs

$$n = 10^5 \quad d = 3 \quad \lambda = 1.1$$



SDP: a better relaxation?

- Maximize $\sum_{i,j} A_{ij}^{\text{cen}} x_i x_j$ over $x \in \{+1, -1\}^n$

it is equivalent to maximize $\sum_{i,j} A_{ij}^{\text{cen}} X_{ij} \equiv \langle A^{\text{cen}}, X \rangle$

subject to $X \in \mathbb{R}^{n \times n}$, $X \succeq 0$ (i.e. all eigenvalues ≥ 0)

$X_{ii} = 1$ and X being of rank 1

- SDP relaxes the rank and maximizes $\langle A^{\text{cen}}, X \rangle$ over the convex space of positive semidefinite matrices
- The maximizer is a matrix of rank m with $m \in [1, n]$ to be projected back on a rank 1 matrix...

$$X^{\text{opt}} \longrightarrow \hat{x}^{\text{SDP}} (\hat{x}^{\text{SDP}})^{\text{T}}$$

SDP-based algorithm

- Maximize $\langle A^{\text{cen}}, X \rangle$ over rank- m matrices = correlation matrices between m -components variables of unit norm

$$C_{ij} = \underline{x}_i \cdot \underline{x}_j, \quad \text{with } \underline{x}_i \in \mathbb{R}^m, \quad \|\underline{x}_i\|^2 = \underline{x}_i \cdot \underline{x}_i = 1$$

- Maximize $\sum_{(ij) \in E} \underline{x}_i \cdot \underline{x}_j$ subject to $\sum_i \underline{x}_i = \underline{0}$

by greedy T=0 dynamics (very fast! no gradient used)

- Given the maximizer $\underline{x}^* = \{\underline{x}_1^*, \dots, \underline{x}_n^*\}$
compute the empirical covariance matrix ($m \times m$)

$$\Sigma_{jk} = \frac{1}{n} \sum_{i=1}^n (\underline{x}_i^*)_j (\underline{x}_i^*)_k$$

- Project on its principal eigenvector $\hat{x}_i^{\text{SDP}} = \text{sign}(\underline{x}_i \cdot \underline{v}_1)$

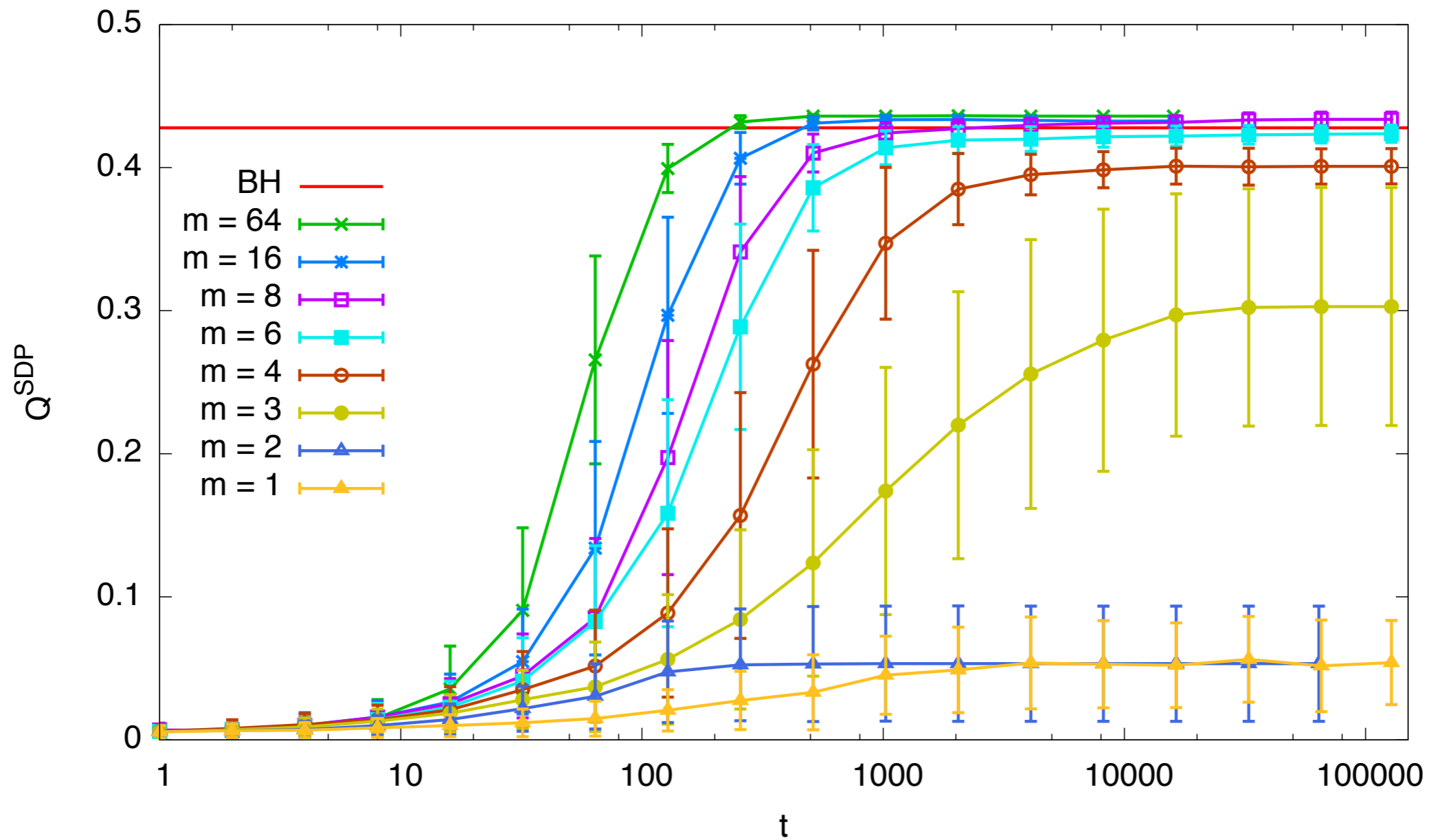
<http://web.stanford.edu/~montanar/SDPgraph/>

SDP-based algorithm

- Algorithm complexity $O(n m t_{\text{conv}})$ and quality of inference do depend on m
 - $m=1$ \rightarrow ML, **very rough** objective function, NP-hard
 - $m=n$ \rightarrow SDP, **convex** objective function
no local maxima for $m > \sqrt{2n}$ [Burer, Monteiro, 2003]
 - $m>1$, **but small** \rightarrow **smooth enough** objective function ?
local minima are "close enough" $O(m^{-1/2})$
to global minimum [Montanari, 2016]
- Running times grows very mildly with m and n
e.g. if stopping rule is max variation $< 10^{-3}$ $\rightarrow t_{\text{conv}} \propto n^{0.22}$

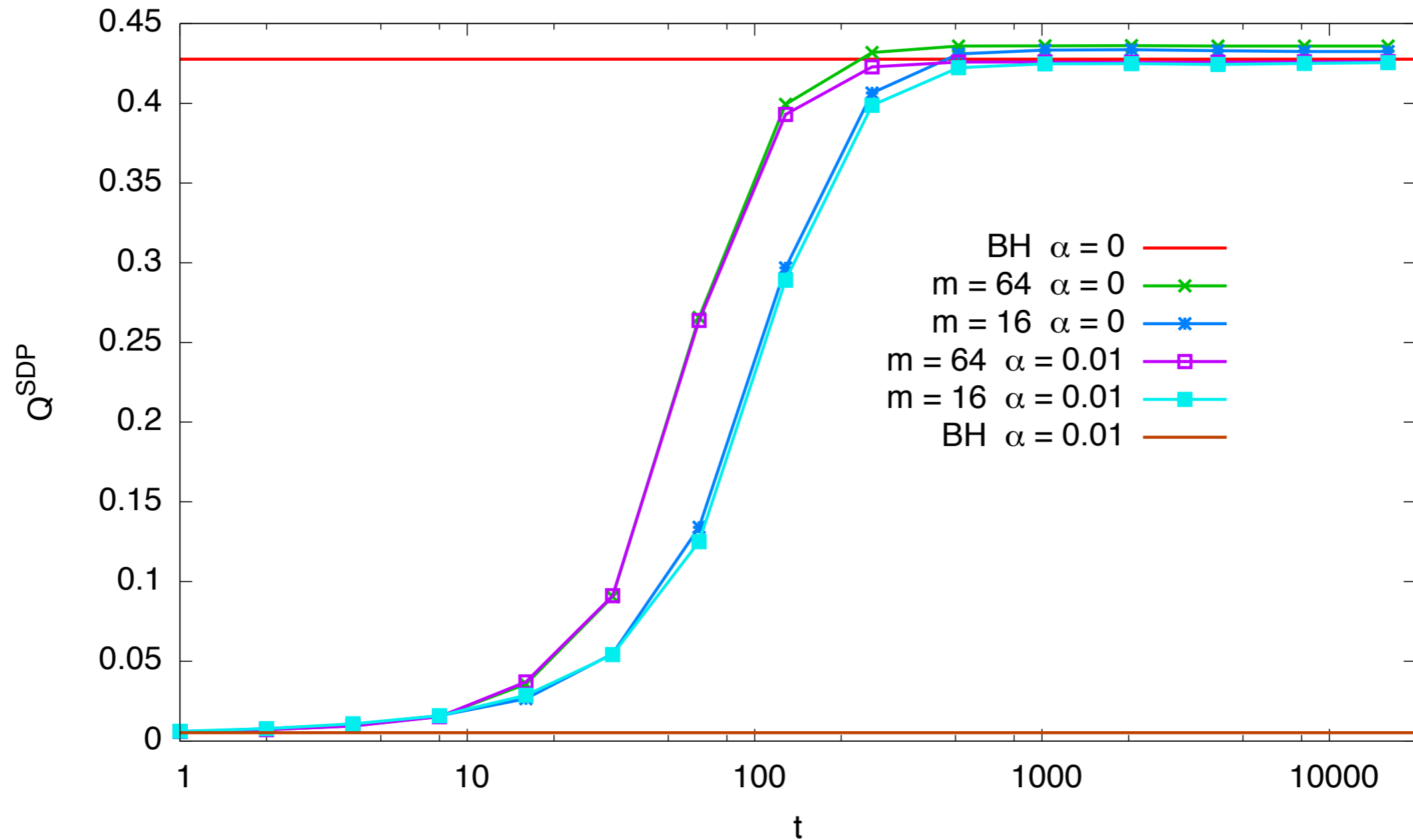
Small m values are fine!

$$n = 4 \cdot 10^4 \quad d = 3 \quad \lambda = 1.1 \quad \alpha = 0.0$$



The algorithm is very robust!

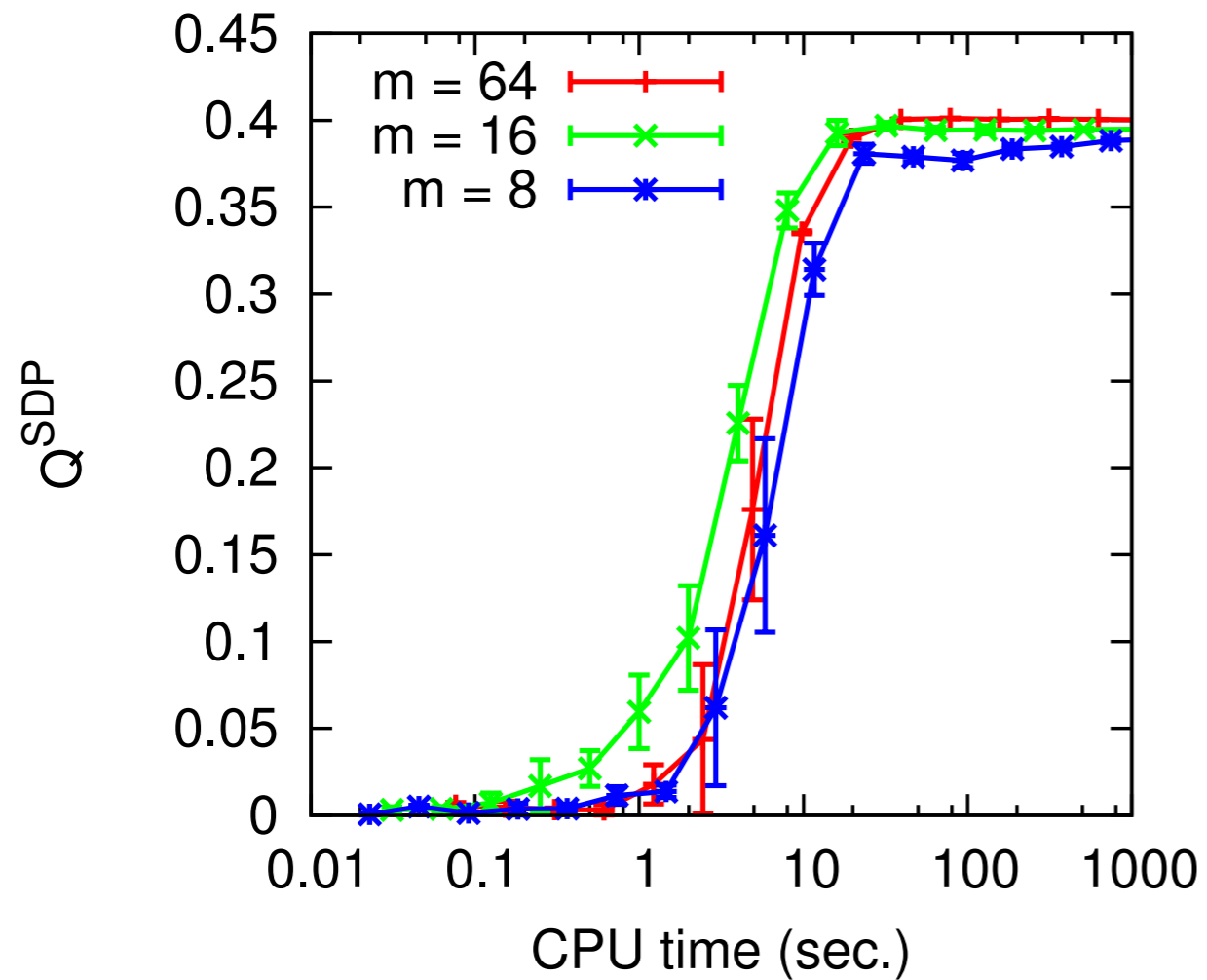
$$n = 4 \cdot 10^4 \quad d = 3 \quad \lambda = 1.1$$



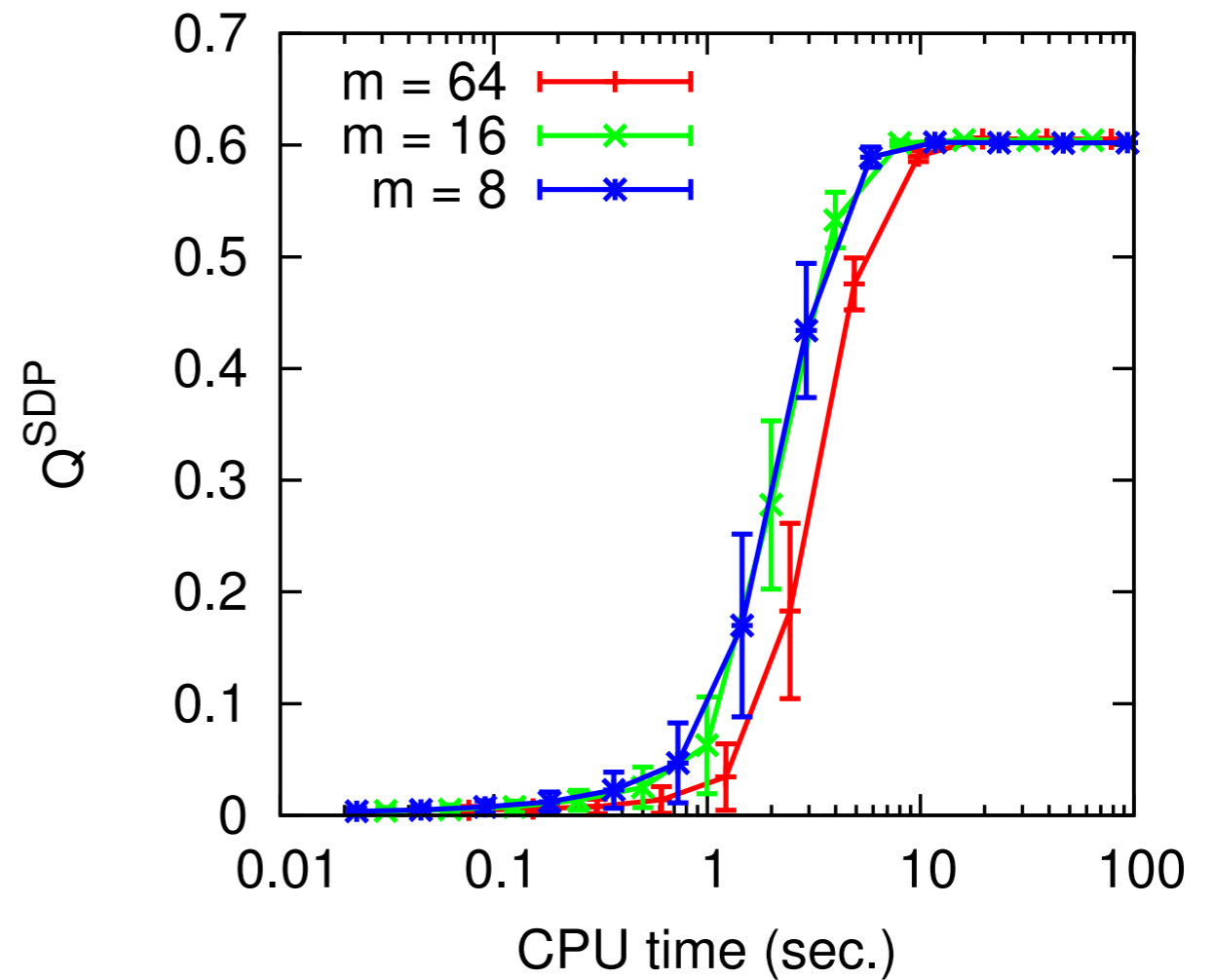
The algorithm is very fast!

$$\underline{n = 10^5} \quad d = 3$$

$$\lambda = 1.1$$

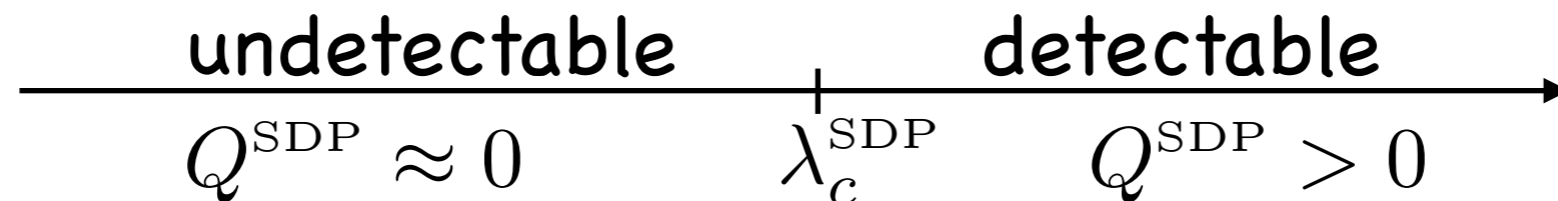


$$\lambda = 1.2$$



SDP optimality

- Analytical predictions on SDP-based hidden partition detection (signal recovery) in the limit $m \rightarrow \infty$



- A simpler synchronization problem:

$$\mathbf{Y} = \frac{\lambda}{n} \mathbf{x}_0 \mathbf{x}_0^* + \mathbf{W}$$

- Recovery \mathbf{x}_0 given the noisy relative positions in \mathbf{Y}

- Different models: $\mathbf{x}_0 \in \mathbb{R}^n$ $W_{ij} \sim \text{N}(0, 1/n)$

$$\mathbf{x}_0 \in \mathbb{C}^n \quad W_{ij} \sim \text{CN}(0, 1/n)$$

Different estimators

- Bayes optimal

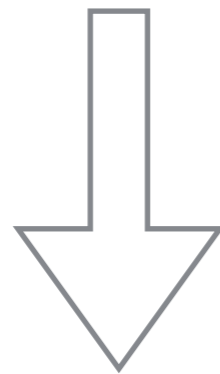
$$\hat{\mathbf{x}}^{\text{Bayes}}(\mathbf{Y}) = \mathbb{E}\{\mathbf{x} \mid (\lambda/n)\mathbf{x}\mathbf{x}^* + \mathbf{W} = \mathbf{Y}\}$$

- Maximum likelihood

$$\hat{\mathbf{x}}^{\text{ML}}(\mathbf{Y}) = c(\lambda) \operatorname{argmax}_{\mathbf{x} \in \{+1, -1\}^n} \langle \mathbf{x}, \mathbf{Y} \mathbf{x} \rangle$$

- SDP

$$\begin{aligned} & \text{maximize} && \langle \mathbf{X}, \mathbf{Y} \rangle, \\ & \text{subject to} && \mathbf{X} \succeq 0, \quad X_{ii} = 1 \quad \forall i \in [n] \end{aligned}$$



$$\hat{\mathbf{x}}^{\text{SDP}}(\mathbf{Y}) = \sqrt{n} c^{\text{SDP}}(\lambda) \mathbf{v}_1(\mathbf{X}_{\text{opt}}(\mathbf{Y}))$$

$$\operatorname{argmax}_{\underline{\mathbf{x}}} \sum_{i,j} \operatorname{Re}(Y_{ij} \underline{\mathbf{x}}_i \cdot \underline{\mathbf{x}}_j) \quad \underline{\mathbf{x}}_i \in \mathbb{F}^m, \quad \|\underline{\mathbf{x}}_i\| = 1$$

Statistical physics approach

- Unified framework: statistical physics models with m -component variables: $\underline{x}_i \in \mathbb{F}^m$, $\|\underline{x}_i\| = 1$

$$P(\underline{x}) = \frac{1}{Z} \exp \left[2m\beta \sum_{i < j} \operatorname{Re}(Y_{ij} \underline{x}_i \cdot \underline{x}_j) \right]$$

- **Bayes**: $m = 1$, $\beta = \begin{cases} \lambda/2 & \text{if } \mathbb{F} = \mathbb{R} \\ \lambda & \text{if } \mathbb{F} = \mathbb{C} \end{cases}$
- **ML**: $m = 1$, $\beta \rightarrow \infty$
- **SDP**: $m \rightarrow \infty$, $\beta \rightarrow \infty$

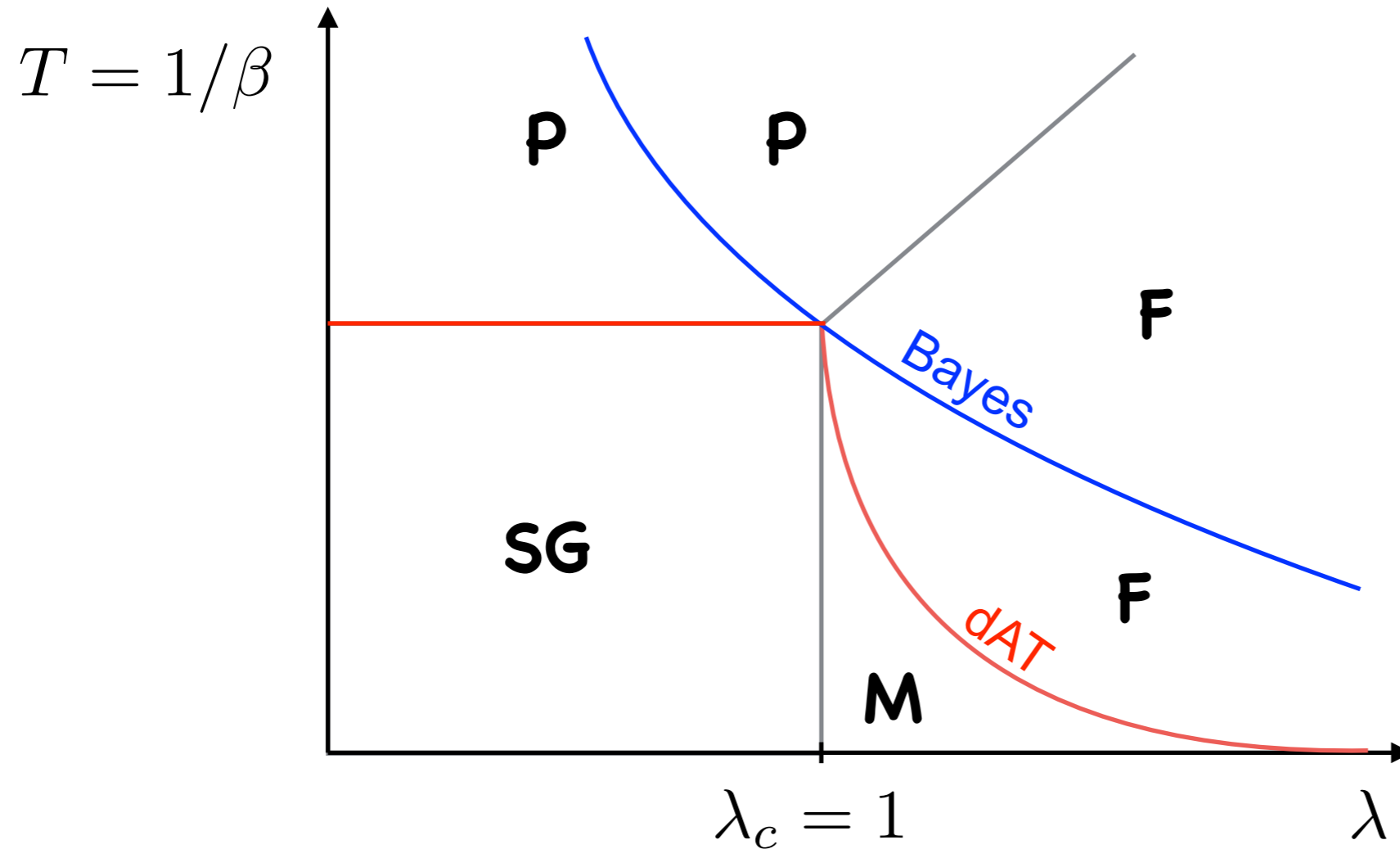
Statistical physics approach

Ising variables ($m=1$), dense graph

Sherrington-Kirkpatrick model

$$H = - \sum_{i < j} J_{ij} s_i s_j \quad s_i \in \{-1, 1\}$$

$$J_{ij} \sim N(\lambda/n, 1/n)$$



- P**: $Q=0$ easy
- F**: $Q>0$ easy
- SG**: $Q=0$ hard
- M**: $Q>0$ hard

Statistical physics approach

- Ansatz for the marginals in m-component dense models

$$P_i(\underline{x}_i) = \frac{1}{Z_i} \exp \left[2m\beta (\underline{\xi}_i^\top \underline{x}_i + \underline{x}_i^\top \mathbf{C}_i \underline{x}_i) \right]$$

$$\underline{x}_i \in \mathbb{F}^m, \|\underline{x}_i\| = 1 \quad \underline{\xi}_i \sim \mathcal{N}(\underline{\mu}, \mathbf{Q}) \quad \mathbf{C}_i = \mathbf{C}$$

- Self consistency equations in the dense case

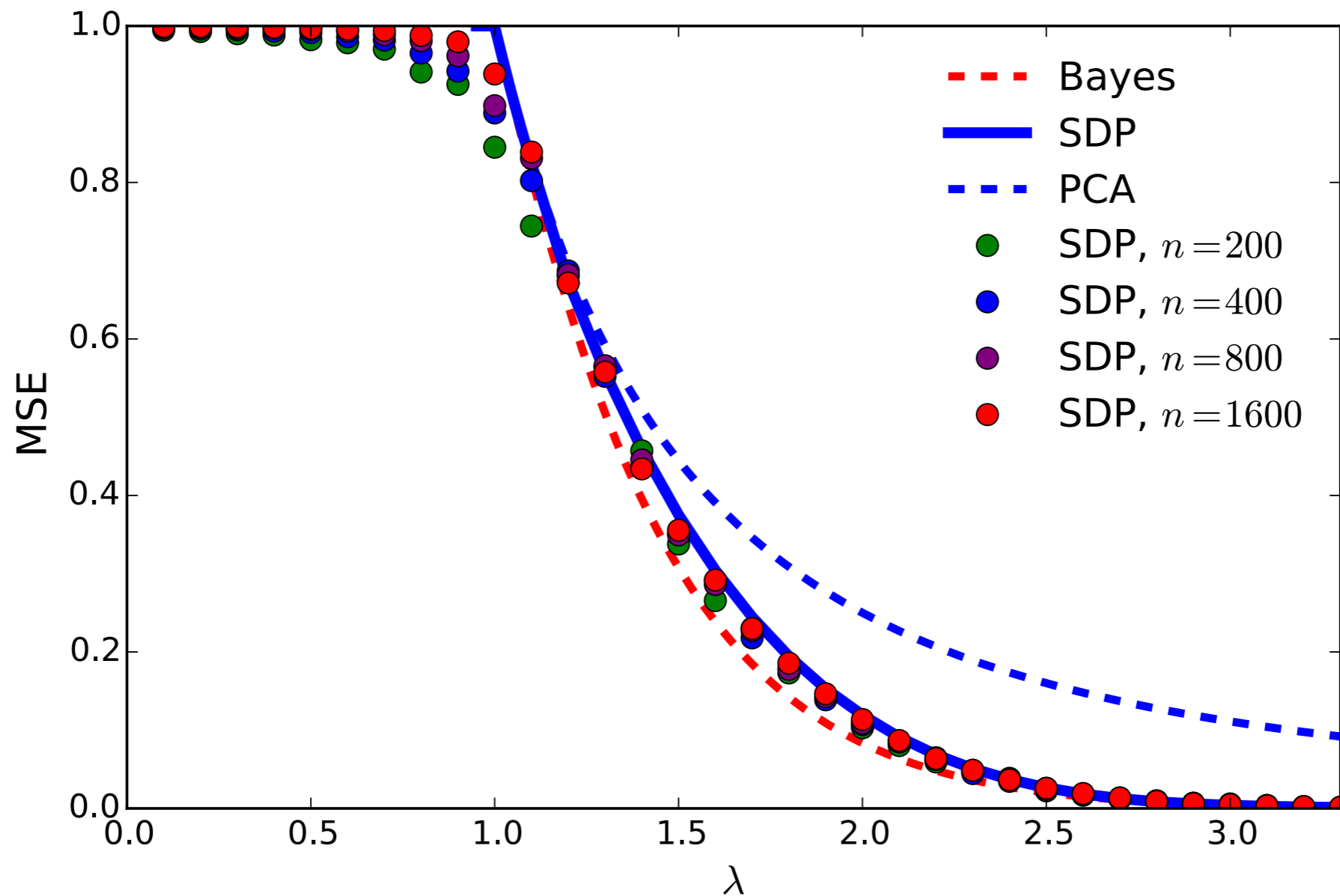
$$\underline{\mu} = \lambda \mathbb{E}[\langle \underline{x} \rangle]$$

$$\mathbf{Q} = \mathbb{E}[\langle \underline{x} \rangle \langle \underline{x}^\top \rangle]$$

$$\mathbf{C} = \beta m \mathbb{E}[\langle \underline{x} \underline{x}^\top \rangle - \langle \underline{x} \rangle \langle \underline{x}^\top \rangle]$$

Analytical solution: dense real case

$$\text{MSE}_n(\hat{\boldsymbol{x}}) \equiv \frac{1}{n} \mathbb{E} \left\{ \min_{s \in \{+1, -1\}} \|\hat{\boldsymbol{x}}(\boldsymbol{Y}) - s \boldsymbol{x}_0\|_2^2 \right\}$$

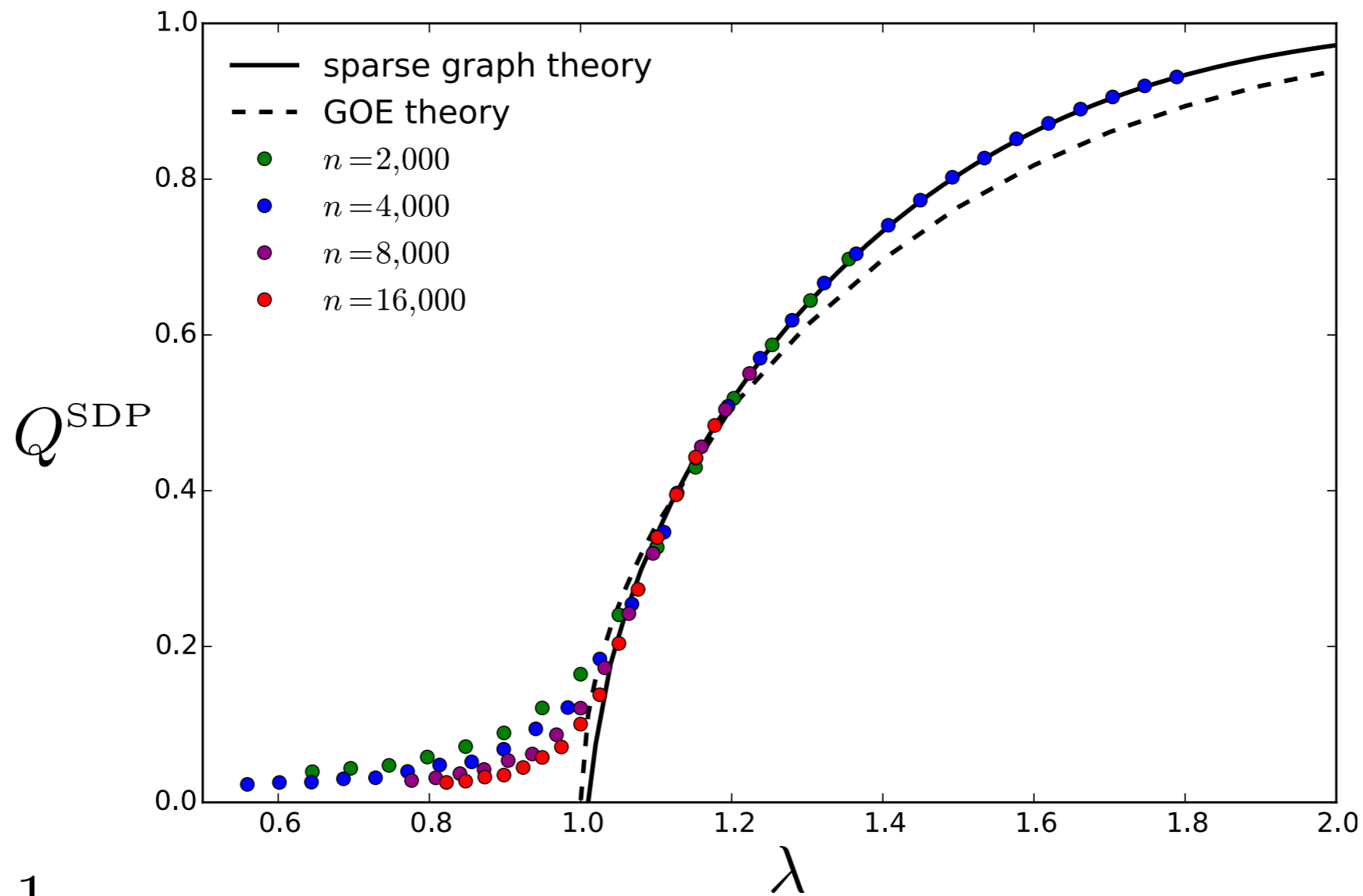


Analytical solution: sparse case (SBM)

- Run the SDP-based algorithm for very large m values
- Approximate ansatz (exact in the large d limit)

Analytical solution: sparse case (SBM)

$$d = 5 \quad N_{\text{samples}} = 500$$

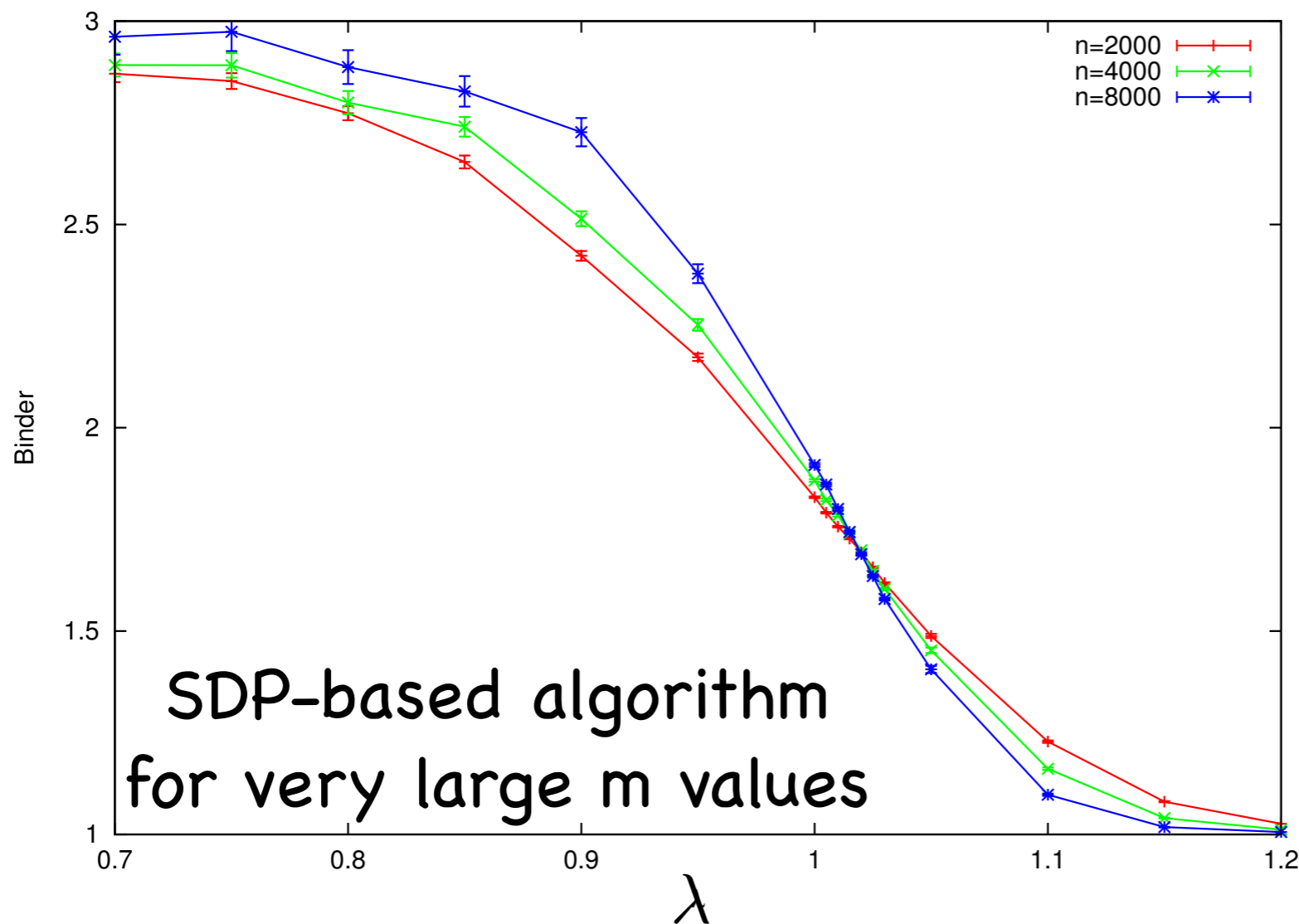


$$Q^{\text{SDP}} = \frac{1}{n} \mathbb{E} \{ |\langle \mathbf{x}^{\text{SDP}}, \mathbf{x}_0 \rangle| \}$$

SDP-based algorithm
for very large n values

Analytical solution: sparse case (SBM)

- Crossing of the Binder cumulants to locate exactly λ_c^{SDP}



$$d = 5$$

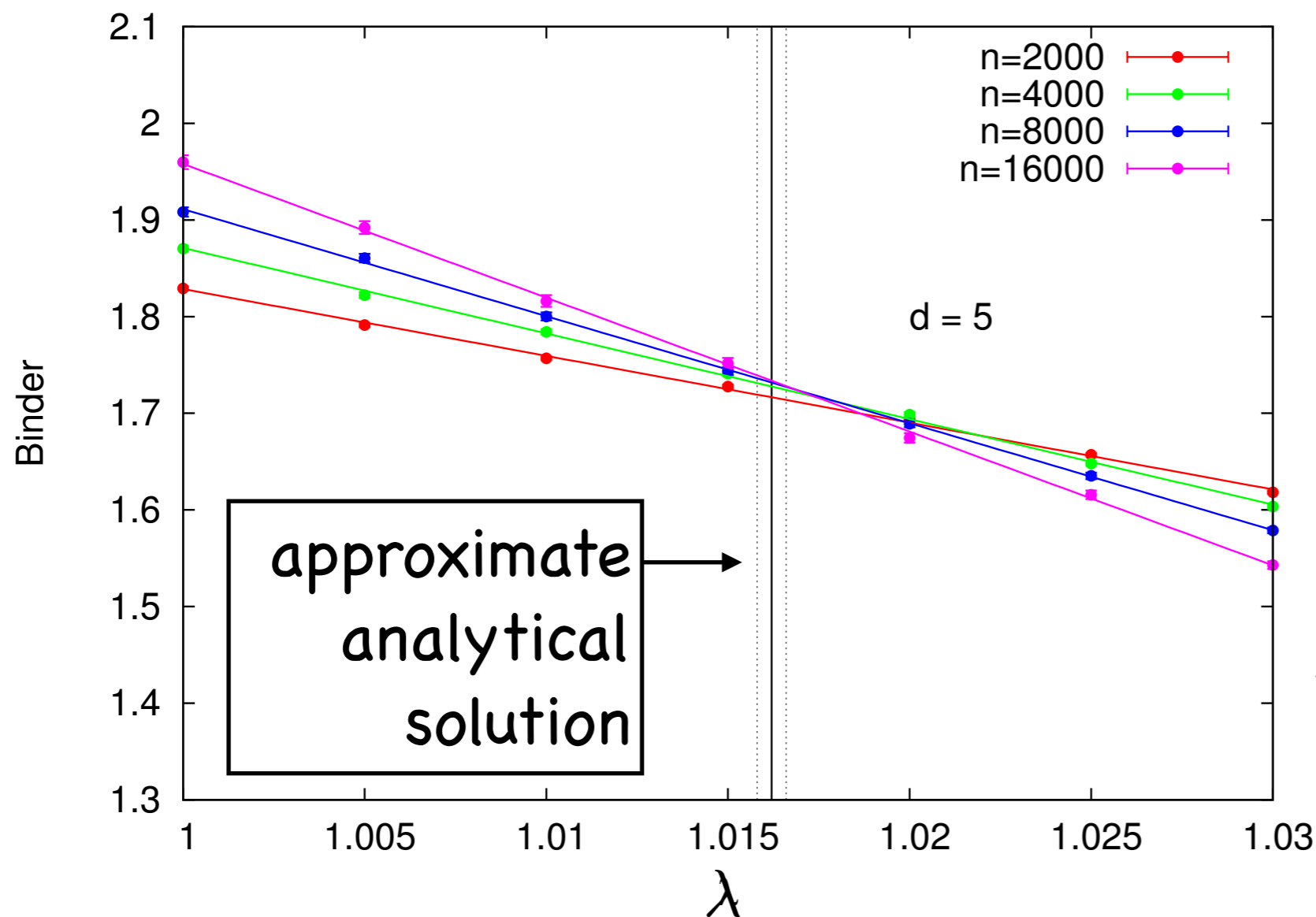
$$N_{\text{samples}} \geq 10^5$$

roughly 2 years
of CPU time

$$B = \frac{n \mathbb{E}\{\langle \mathbf{x}^{\text{SDP}}, \mathbf{x}_0 \rangle^4\}}{(\mathbb{E}\{\langle \mathbf{x}^{\text{SDP}}, \mathbf{x}_0 \rangle^2\})^2}$$

Analytical solution: sparse case (SBM)

- Crossing of the Binder cumulants to locate exactly λ_c^{SDP}



$$d = 5$$

$$N_{\text{samples}} \geq 10^5$$

roughly 2 years
of CPU time

$$B = \frac{n \mathbb{E}\{\langle \mathbf{x}^{\text{SDP}}, \mathbf{x}_0 \rangle^4\}}{(\mathbb{E}\{\langle \mathbf{x}^{\text{SDP}}, \mathbf{x}_0 \rangle^2\})^2}$$

SDP-based algorithm
for very large n values

Approximate analytical solution (SBM)

- In the recovery phase we assume the $O(m)$ symmetry to break along the first component, while preserving $O(m-1)$

$$\underline{x}_i = (s_i, \boldsymbol{\tau}_i), \quad s_i \in \mathbb{R}, \quad \boldsymbol{\tau}_i \in \mathbb{R}^{m-1}$$

- We write the marginal for \underline{x}_i as

$$\exp \left\{ 2\beta\sqrt{mc_i} \langle \mathbf{z}_i, \boldsymbol{\tau}_i \rangle + 2\beta mh_i s_i - \beta mr_i s_i^2 + O_m(1) \right\} \delta \left(s_i^2 + \|\boldsymbol{\tau}_i\|_2^2 - 1 \right)$$

with $\mathbf{z}_i \sim N(0, I_{m-1})$

- Approximate because the \mathbf{z}_i are correlated
- It should be valid in the limits $d \rightarrow 1$ and $d \rightarrow \infty$

Approximate analytical solution (SBM)

$$\exp \left\{ 2\beta \sqrt{m c_i} \langle \mathbf{z}_i, \boldsymbol{\tau}_i \rangle + 2\beta m h_i s_i - \beta m r_i s_i^2 + O_m(1) \right\} \delta \left(s_i^2 + \|\boldsymbol{\tau}_i\|_2^2 - 1 \right)$$

- Cavity method \rightarrow self consistency equation for marginals

$$c_0 = \sum_{i=1}^k \frac{c_i}{\rho_i^2},$$

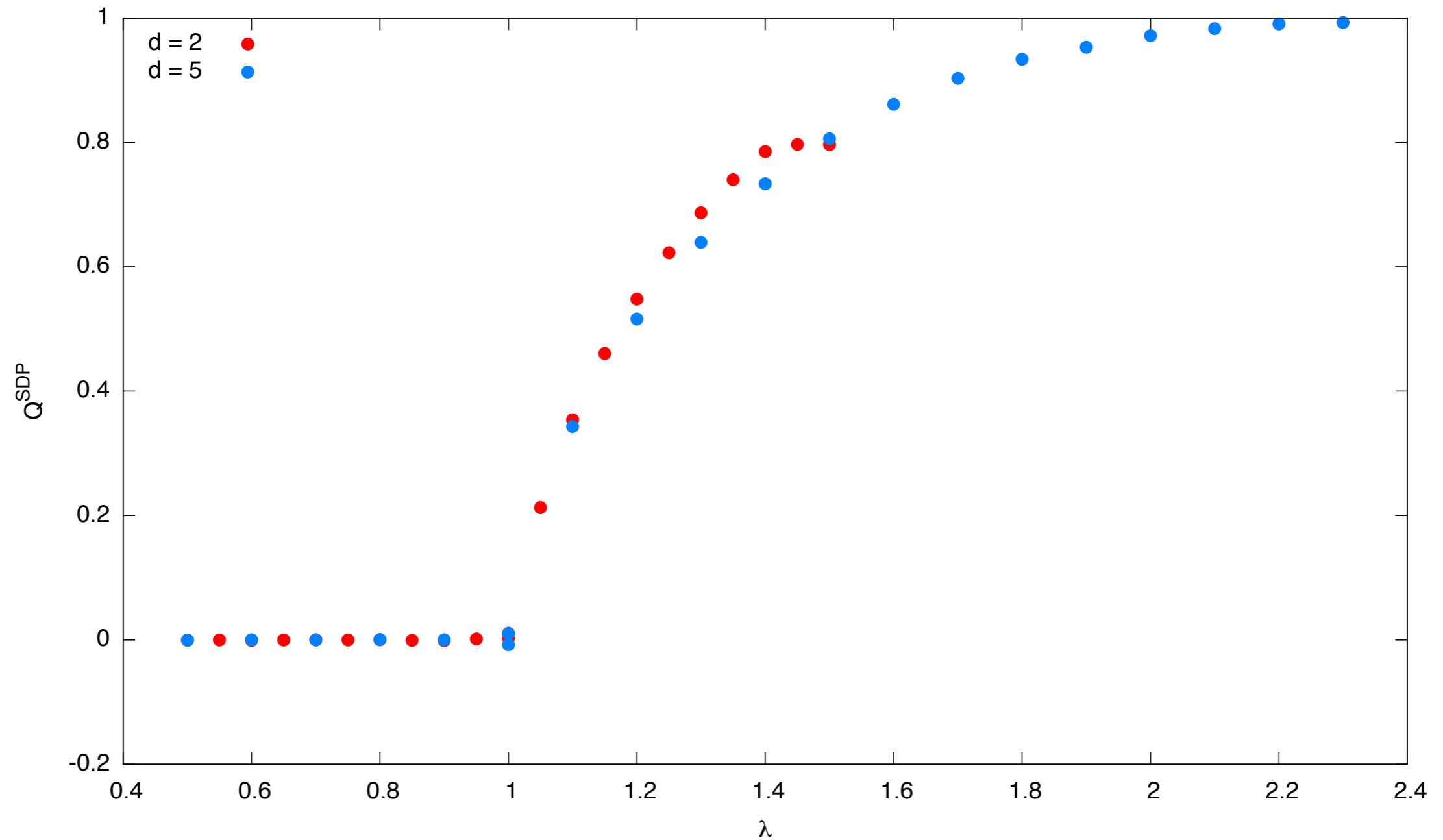
$$h_0 = \sum_{i=1}^k \frac{h_i}{\rho_i + r_i},$$

$$r_0 = \sum_{i=1}^k \left\{ \frac{1}{\rho_i} - \frac{1}{\rho_i + r_i} + \left(1 + \frac{(1 + c_i)r_i}{\rho_i^3} \right)^{-1} \frac{h_i^2}{(\rho_i + r_i)^3} \right\}$$

$$1 = \frac{h_i^2}{(\rho_i + r_i)^2} + \frac{1 + c_i}{\rho_i^2}$$

- Solve by population dynamics
- At the fixed point $Q^{\text{SDP}} = \mathbb{E}[\text{sign}(h^*)]$

Approximate analytical solution (SBM)



Approximate analytical solution (SBM)

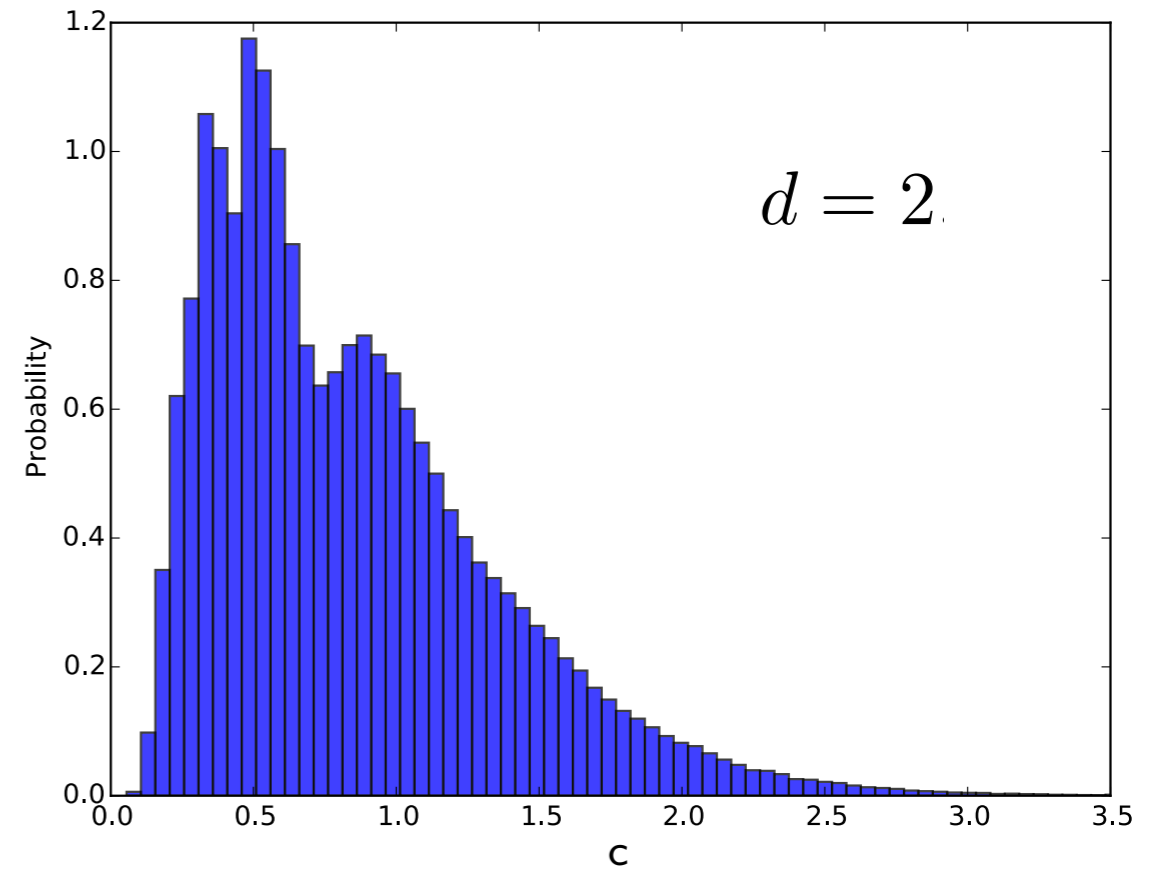
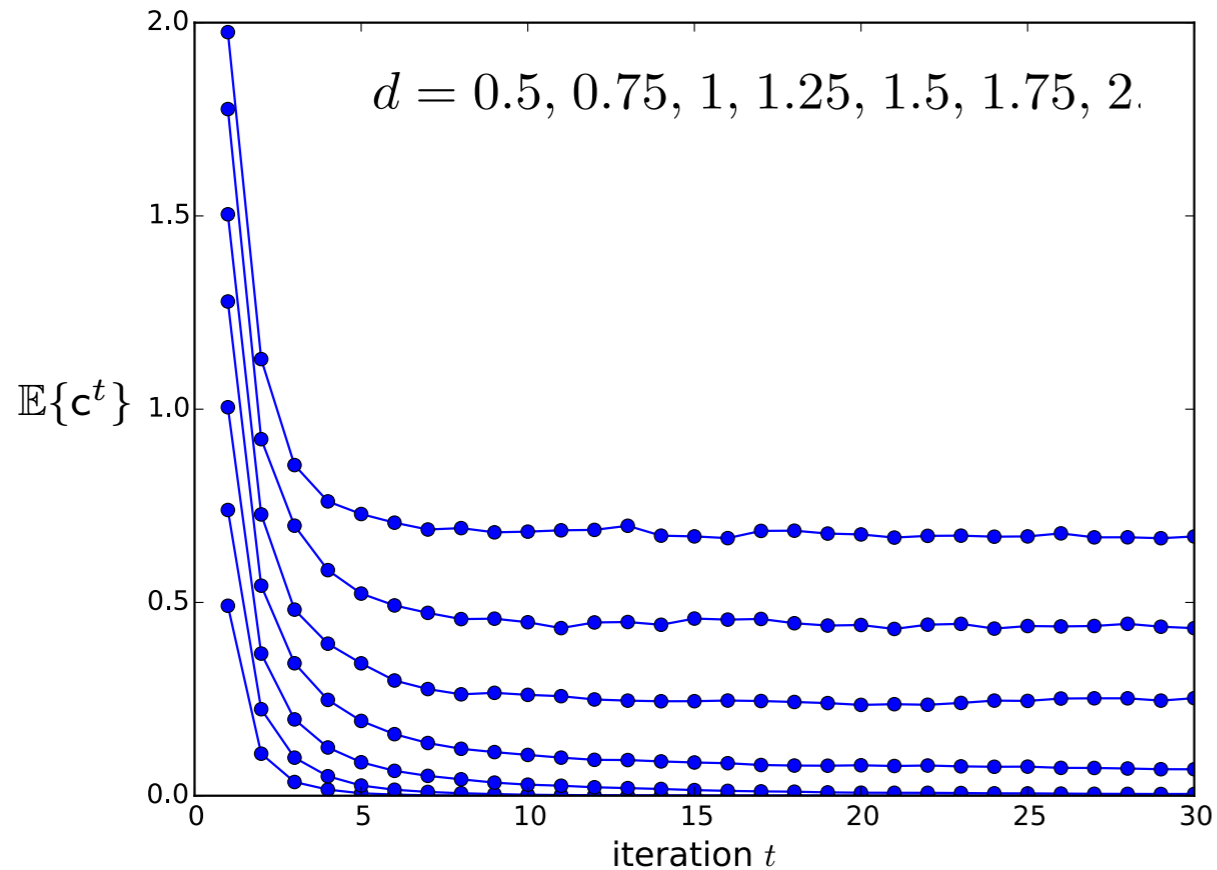
- To linear order in $h \implies r_i = 0$

$$c_0 = \sum_{i=1}^k \frac{c_i}{\rho_i^2},$$
$$h_0 = \sum_{i=1}^k \frac{h_i}{\rho_i + r_i},$$
$$r_0 = \sum_{i=1}^k \left\{ \frac{1}{\rho_i} - \frac{1}{\rho_i + r_i} + \left(1 + \frac{(1 + c_i)r_i}{\rho_i^3} \right)^{-1} \frac{h_i^2}{(\rho_i + r_i)^3} \right\}$$
$$1 = \frac{h_i^2}{(\rho_i + r_i)^2} + \frac{1 + c_i}{\rho_i^2}$$



$$c_0 = \sum_{i=1}^k \frac{c_i}{1 + c_i},$$
$$h_0 = \sum_{i=1}^k \frac{h_i}{\sqrt{1 + c_i}}$$

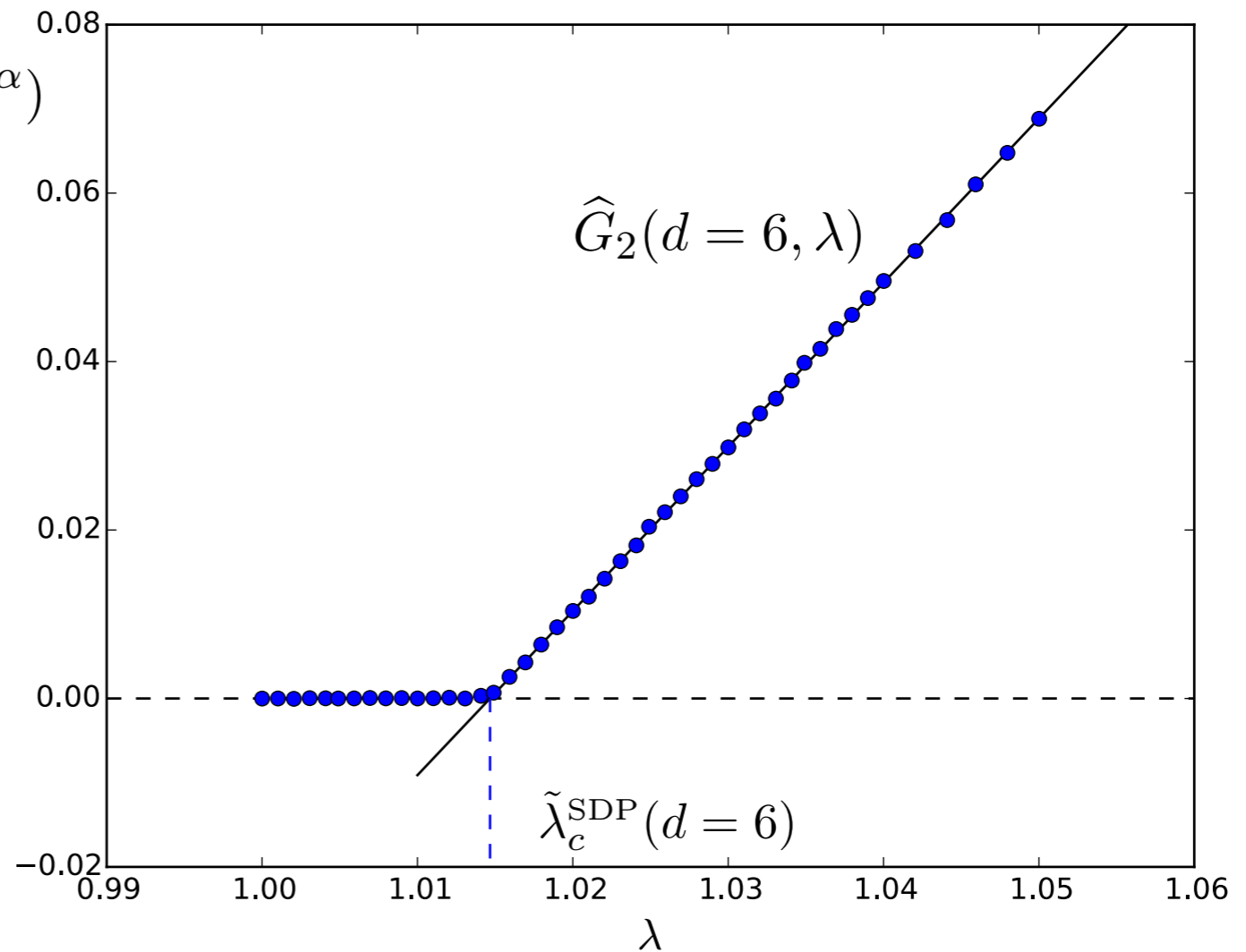
Approximate analytical solution (SBM)



$$(c^{t+1}; h^{t+1}) \stackrel{d}{=} \left(\sum_{i=1}^{L_+ + L_-} \frac{c_i^t}{1 + c_i^t}; \sum_{i=1}^{L_+ + L_-} \frac{s_i h_i^t}{\sqrt{1 + c_i^t}} \right)$$

Approximate analytical solution (SBM)

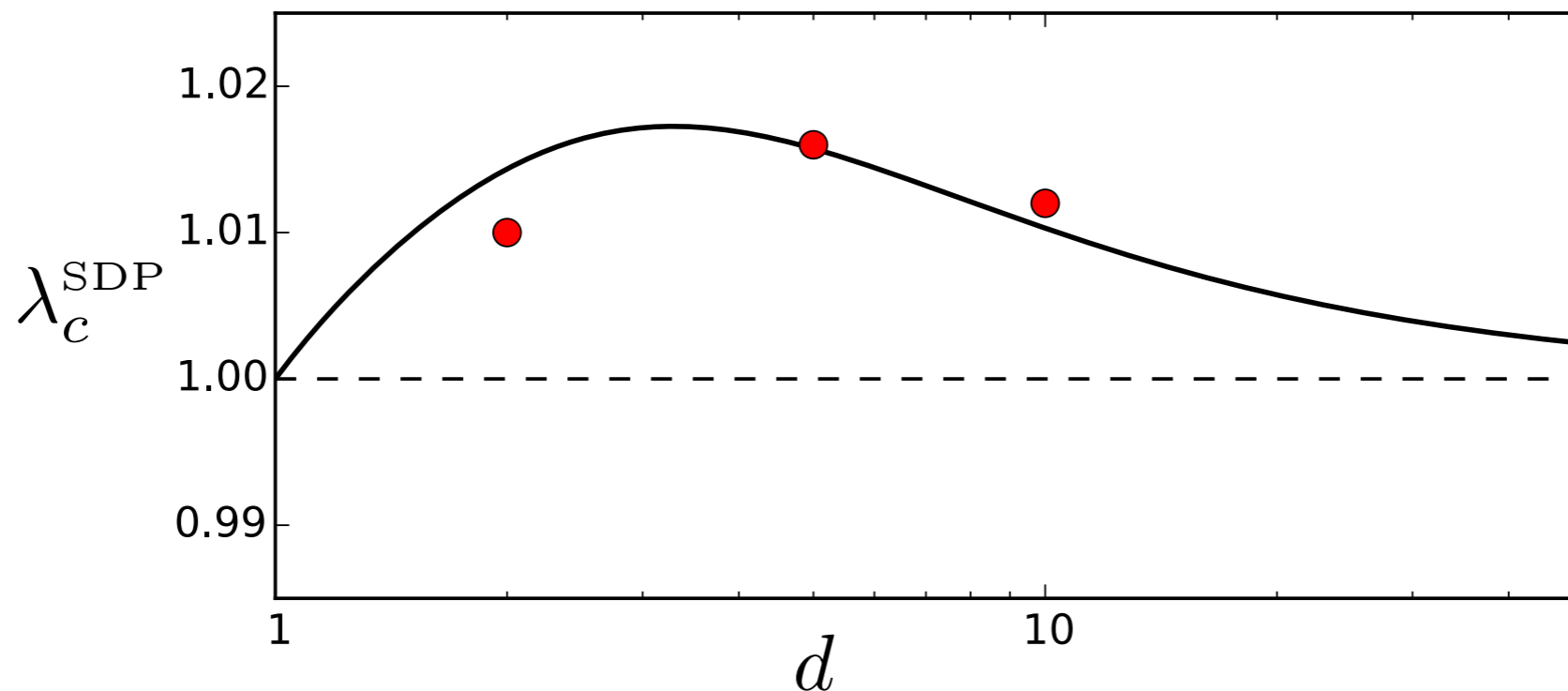
$$G_\alpha(d, \lambda) \equiv \liminf_{t \rightarrow \infty} \frac{1}{t\alpha} \log \mathbb{E}(|h^t|^\alpha)$$



$$(c^{t+1}; h^{t+1}) \stackrel{d}{=} \left(\sum_{i=1}^{L_+ + L_-} \frac{c_i^t}{1 + c_i^t}; \sum_{i=1}^{L_+ + L_-} \frac{s_i h_i^t}{\sqrt{1 + c_i^t}} \right)$$

Analytical solution: sparse case (SBM)

- SDP at most 2% sub-optimal!



- **Red points:** numerical solution of the replica/cavity equations (crossing of Binder cumulants)
- **Black line:** approximated analytical solution

Take-home messages

- SDP relaxations are very effective:
 - robust and quasi-optimal
 - may outperform spectral relaxations
- Better than SDP are SDP-inspired algorithms (small m)
<http://web.stanford.edu/~montanar/SDPgraph/>
- It is worth studying the statistical physics of models with m -component variables:
 - unifying framework to study and solve several estimators in statistical inference
 - different physics, better algorithms