# Recent Advances in Counting Sparse Graphs

Sourav Chatterjee

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- Was an open question for a long time. (History in the next two slides.) Still not fully resolved. This deceptively difficult problem falls in the intersection of large deviations, concentration of measure, and random graph theory.

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- Was an open question for a long time. (History in the next two slides.) Still not fully resolved. This deceptively difficult problem falls in the intersection of large deviations, concentration of measure, and random graph theory.
- This is quite surprising since T is simply a third degree polynomial of independent random variables. We know everything about linear functions of independent random variables, so why all this difficulty for a third degree polynomial?

▶ After a long line of successively improving tail bounds by various authors, Kim and Vu (2004) and Janson, Oleszkiewicz and Ruciński (2004) showed that if  $p \ge N^{-1} \log N$ , then

 $e^{-c_1(\delta) \mathcal{N}^2 p^2 \log(1/p)} \leq \mathbb{P}(\mathcal{T} \geq (1+\delta)\mathbb{E}(\mathcal{T})) \leq e^{-c_2(\delta) \mathcal{N}^2 p^2}$  ,

where  $c_1(\delta)$  and  $c_2(\delta)$  are constants that depend on  $\delta$  only.

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- For a certain range of fixed values of p and δ, Chatterjee and Dey (2009) — using Stein's method for concentration inequalities — computed the exact constant c(δ, p) such that, as N → ∞,

$$\mathbb{P}(T \geq (1+\delta)\mathbb{E}(T)) = e^{-c(\delta,p)N^2(1+o(1))}$$
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▶ This result, however, did not cover the full range of  $\delta$  and  $p_{\pm}$ 

For the case of fixed p, the problem was fully resolved by Chatterjee and Varadhan (2011), who exhibited the constant c(δ, p) as the solution of a variational problem. This formula was conjectured in an unpublished manuscript of Bolthausen, Comets and Dembo (2003).

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- However, the regularity lemma has no satisfactory analog for sparse graphs, certainly when p is of order N<sup>-α</sup> for some positive α.
- This made it impossible to extend the proof to the sparse case.

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Theorem (Lubetzky and Zhao (2014)) If  $N \to \infty$  and  $p \to 0$  slower than  $N^{-1/42}$ , then

$$\mathbb{P}(T \ge (1+\delta)\mathbb{E}(T))$$
  
= exp $\left(-(1+o(1))\min\left\{\frac{\delta^{2/3}}{2}, \frac{\delta}{3}\right\}N^2p^2\log\frac{1}{p}\right).$ 

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- Conjecture: This formula is valid if  $p \rightarrow 0$  slower than  $N^{-1/2}$ .
- The proof uses the theory of nonlinear large deviations developed by Chatterjee and Dembo (2014). This will be described later in this talk.

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- ▶ Define a polynomial P<sub>H\*</sub>(x) := ∑<sub>k</sub> i<sub>H\*</sub>(k)x<sup>k</sup>, where i<sub>H\*</sub>(k) is the number of k-element independent sets in H<sup>\*</sup>.

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- Let  $H_{N,p}$  be the number of homomorphisms of H into G(N, p).
- ► A homomorphism is a map from the vertex set of H into the vertex set of G(N, p) that preserves edges. For example, if H is a triangle, then H<sub>N,p</sub> is six times the number of triangles in G(N, p).

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#### Large deviations for subgraph counts

Theorem (Bhattacharya, Ganguly, Lubetzky and Zhao (2015)) For any  $\delta > 0$ , there is a unique positive number  $\theta = \theta(H, \delta)$  that solves  $P_{H^*}(\theta) = 1 + \delta$ . There is a constant  $\alpha_H > 0$  depending only on H, such that if  $N \to \infty$  and  $p \to 0$  slower than  $N^{-\alpha_H}$ ,

$$\mathbb{P}(H_{N,p} \ge (1+\delta)\mathbb{E}(H_{N,p})) = \exp\left(-(1+o(1))c(\delta)N^2p^{\Delta}\log\frac{1}{p}\right),$$

where

$$c(\delta) = \begin{cases} \min\{\theta, \frac{1}{2}\delta^{2/k}\} & \text{if } H \text{ is regular,} \\ \theta & \text{if } H \text{ is irregular.} \end{cases}$$

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Example: If  $H = C_4$ , then

$$m{c}(\delta) = egin{cases} rac{1}{2}\sqrt{\delta} & ext{if } \delta < 16, \ -1 + \sqrt{1 + rac{1}{2}\delta} & ext{if } \delta \geq 16. \end{cases}$$

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- ► Classical large deviations theory well-suited for linear *f*.
- May be quite nontrivial even for very simple nonlinear f, as we saw in the random graph example. Problems tackled on ad hoc basis.

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- ► Classical large deviations theory well-suited for linear *f*.
- May be quite nontrivial even for very simple nonlinear f, as we saw in the random graph example. Problems tackled on ad hoc basis.
- Can there be a more unified approach?

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• For 
$$x = (x_1, ..., x_n) \in [0, 1]^n$$
, define

$$I_p(x) := \sum_{i=1}^n \left( x_i \log \frac{x_i}{p} + (1-x_i) \log \frac{1-x_i}{1-p} \right).$$

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For each  $t \in \mathbb{R}$ , define

$$\phi_{p}(t):=\inf\{I_{p}(x)\,:\,x\in[0,1]^{n} ext{ such that }f(x)\geq tn\}\,.$$

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In many problems, it turns out that

$$\mathbb{P}(f(Y) \ge tn) \approx \exp(-\phi_p(t)). \tag{(\star)}$$

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$$\phi_{\rho}(t) := \inf\{I_{\rho}(x) : x \in [0,1]^n \text{ such that } f(x) \ge tn\}.$$

In many problems, it turns out that

$$\mathbb{P}(f(Y) \ge tn) \approx \exp(-\phi_p(t)). \qquad (\star)$$

- In particular, this is true in great generality for linear functions.
- I will describe a sufficient condition under which the above approximation is valid for nonlinear maps.

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The approximation  $(\star)$  is valid when, in addition to some smoothness conditions on the function f, the gradient vector  $\nabla f(x) = (\partial f / \partial x_1, \dots \partial f / \partial x_n)$  may be approximately encoded by o(n) bits of information.

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- Many notable results on sharp upper and lower bounds for tail probabilities of nonlinear functions (Talagrand, Kim, Vu, Latała, ....) that hold up to constant factors in the exponent, but no results about the precise approximation (\*).

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- Many notable results on sharp upper and lower bounds for tail probabilities of nonlinear functions (Talagrand, Kim, Vu, Latała, ....) that hold up to constant factors in the exponent, but no results about the precise approximation (\*).
- Some preliminary work in Chatterjee & Dey (2009).

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- Thus, for this f, the gradient vector cannot be approximately encoded by o(n) many bits. To know ∇f(x), even approximately, we need to know the values of all the x<sub>i</sub>'s.
- One can check that the approximation (\*) is not valid for this f.

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▶ Thus, for this f, the gradient vector is approximately encoded by the single quantity  $n^{-1} \sum x_j$ .

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- The large deviation probabilities for this function satisfy the approximation (\*).

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▶ Let n = N(N-1)/2 and let us agree to denote elements of  $\mathbb{R}^n$  as  $x = (x_{ij})_{1 \le i < j \le N}$ , with the convention that  $x_{ii} = 0$  and  $x_{ji} = x_{ij}$ . Define a function  $f : \mathbb{R}^n \to \mathbb{R}$  as

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► Let T be the number of triangles in an Erdős-Rényi random graph G(N, p).

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The plan is to apply the main theorem to this f.

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$$\frac{\partial f}{\partial x_{ij}} = \frac{3}{N} \sum_{k=1}^{N} x_{ik} x_{jk} =: 3a_{ij}(x).$$

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- These two observations allow us to prove the low complexity condition for ∇f.

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What is the precise result for triangle counts?

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• For 
$$x = (x_{ij})_{1 \le i < j \le N}$$
, define

$$I_p(x) := \sum_{1 \le i < j \le N} \left( x_{ij} \log \frac{x_{ij}}{p} + (1 - x_{ij}) \log \frac{1 - x_{ij}}{1 - p} \right)$$

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and

$$T(x):=\frac{1}{6}\sum_{i,j,k}x_{ij}x_{jk}x_{ki},$$

where  $x_{ij} \in [0, 1]$  and  $x_{ji} = x_{ij}$ ,  $x_{ii} = 0$ .

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where  $x_{ij} \in [0, 1]$  and  $x_{ji} = x_{ij}$ ,  $x_{ii} = 0$ . For u > 1 define

$$\psi_p(u) := \inf\{I_p(x) : T(x) \ge u \mathbb{E}(T)\},\$$

where T is the number of triangles in G(N, p).

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Theorem (Chatterjee & Dembo, 2014.) For u > 1 and N sufficiently large (depending only on u),

$$1 - \frac{c \log N}{N^{1/6} p^2} \le \frac{\psi_p(u)}{-\log \mathbb{P}(T \ge u \, \mathbb{E}(T))} \le 1 + \frac{C(\log N)^{33/29}}{N^{1/29} p^{42/29}} \,,$$

where c and C are constants that depend only on u.

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► Lubetzky and Zhao (2014) studied the function ψ<sub>p</sub>(u) to obtain their large deviation result.

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• Given  $\epsilon > 0$ , let  $\mathcal{D}(\epsilon)$  be a finite subset of  $\mathbb{R}^n$  such that for all  $x \in \{0,1\}^n$ , there exists  $d = (d_1, \ldots, d_n) \in \mathcal{D}(\epsilon)$  such that

$$\sum_{i=1}^n (f_i(x)-d_i)^2 \leq n\epsilon^2.$$

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## Main result, contd.

► For 
$$x = (x_1, ..., x_n) \in [0, 1]^n$$
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$$I(x) := \sum_{i=1}^n (x_i \log x_i + (1 - x_i) \log(1 - x_i)).$$

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#### Main result, contd.

► For 
$$x = (x_1, ..., x_n) \in [0, 1]^n$$
, let  
$$I(x) := \sum_{i=1}^n (x_i \log x_i + (1 - x_i) \log(1 - x_i)).$$

Let

$$F:=\log\sum_{x\in\{0,1\}^n}e^{f(x)}.$$

• Given  $\epsilon > 0$ , define

complexity term := 
$$\frac{1}{4} \left( n \sum_{i=1}^{n} b_i^2 \right)^{1/2} \epsilon + 3n\epsilon + \log |\mathcal{D}(\epsilon)|$$
, and  
smoothness term :=  $4 \left( \sum_{i=1}^{n} (ac_{ii} + b_i^2) + \frac{1}{4} \sum_{i,j=1}^{n} (ac_{ij}^2 + b_i b_j c_{ij} + 4b_i c_{ij}) \right)^{1/2} + \frac{1}{4} \left( \sum_{i=1}^{n} b_i^2 \right)^{1/2} \left( \sum_{i=1}^{n} c_{ii}^2 \right)^{1/2} + 3 \sum_{i=1}^{n} c_{ii} + \log 2.$ 

Sourav Chatterjee Recent Advances in Counting Sparse Graphs

Theorem (Chatterjee & Dembo, 2014) For any  $\epsilon > 0$ ,

 $F \leq \sup_{x \in [0,1]^n} (f(x) - I(x)) + \text{complexity term} + \text{smoothness term} ,$ 

and

$$F \geq \sup_{x \in [0,1]^n} (f(x) - I(x)) - rac{1}{2} \sum_{i=1}^n c_{ii} + c_{ii}$$

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Let X = (X<sub>1</sub>,...,X<sub>n</sub>) be a random vector that has probability density proportional to e<sup>f(x)</sup> on {0,1}<sup>n</sup> with respect to the counting measure.

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- For each *i*, define a function  $\hat{x}_i : [0,1]^n \rightarrow [0,1]$  as

$$\hat{x}_i(x) = \mathbb{E}(X_i \mid X_j = x_j, 1 \le j \le n, j \ne i).$$

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- Let  $\hat{X} = \hat{x}(X)$ .
- The first step in the proof is to show that if the smoothness term is small, then

$$f(X) \approx f(\hat{X})$$
 with high probability.

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• Next, define a function  $g: [0,1]^n \times [0,1]^n \to \mathbb{R}$  as

$$g(x,y) := \sum_{i=1}^{n} (x_i \log y_i + (1-x_i) \log(1-y_i)).$$

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The second step is to show that with high probability,

$$g(X, \hat{X}) \approx g(\hat{X}, \hat{X}) = I(\hat{X}).$$

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- ► Let A be the set of all x where  $f(x) \approx f(\hat{x}(x))$  and  $g(x, \hat{x}(x)) \approx I(\hat{x}(x))$ .

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- Suppose that these two steps have been proved.
- Let A be the set of all x where  $f(x) \approx f(\hat{x}(x))$  and  $g(x, \hat{x}(x)) \approx I(\hat{x}(x))$ .
- Since  $X \in A$  with high probability,

$$\frac{\sum_{x \in A} e^{f(x)}}{\sum_{x \in \{0,1\}^n} e^{f(x)}} \approx 1.$$

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► Therefore

$$F = \log \sum_{x \in \{0,1\}^n} e^{f(x)} \approx \log \sum_{x \in A} e^{f(x)}$$
$$\approx \log \sum_{x \in A} e^{f(\hat{x}(x)) - I(\hat{x}(x)) + g(x, \hat{x}(x))}.$$

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Therefore

$$\begin{split} F &= \log \sum_{x \in \{0,1\}^n} e^{f(x)} \approx \log \sum_{x \in A} e^{f(x)} \\ &\approx \log \sum_{x \in A} e^{f(\hat{x}(x)) - I(\hat{x}(x)) + g(x, \hat{x}(x))} \,. \end{split}$$

• Now let  $\epsilon$  be a small positive number.

#### Therefore

$$F = \log \sum_{x \in \{0,1\}^n} e^{f(x)} \approx \log \sum_{x \in A} e^{f(x)}$$
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- Now let  $\epsilon$  be a small positive number.
- Using the low complexity gradient condition, it is possible to construct a set D'(ε) of size e<sup>o(n)</sup> such that for each x ∈ [0, 1]<sup>n</sup> there exists p ∈ D'(ε) such that x̂(x) ≈ p.

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- For each p ∈ D'(ε) let P(p) be the set of all x ∈ {0,1}<sup>n</sup> such that x̂(x) ≈ p.

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• The crucial fact is that for any  $p \in [0, 1]^n$ ,

$$\sum_{x \in \{0,1\}^n} e^{g(x,p)} = 1.$$

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Therefore,

$$\log \sum_{x \in A} e^{f(\hat{x}(x)) - I(\hat{x}(x)) + g(x, \hat{x}(x))}$$
  
$$\leq \log \sum_{p \in \mathcal{D}'(\epsilon)} \sum_{x \in \mathcal{P}(p)} e^{f(\hat{x}(x)) - I(\hat{x}(x)) + g(x, \hat{x}(x))}$$

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► This completes the proof sketch for the upper bound, modulo the two unproved steps. I will now sketch why  $f(X) \approx f(\hat{X})$ . Source Chatteriee Recent Advances in Counting Sparse Graphs

# Proof of $f(X) \approx f(\hat{X})$

► To show this, define

$$h(x) := f(x) - f(\hat{x}(x)).$$

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# Proof of $f(X) \approx f(\hat{X})$

To show this, define

$$h(x) := f(x) - f(\hat{x}(x)).$$
• Let  $u_i(t, x) := f_i(tx + (1 - t)\hat{x}(x))$ , so that
$$h(x) = \int_0^1 \sum_{i=1}^n (x_i - \hat{x}_i(x))u_i(t, x) dt.$$

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• Thus, if  $D := f(X) - f(\hat{X})$ , then

$$\mathbb{E}(D^2) = \int_0^1 \sum_{i=1}^n \mathbb{E}((X_i - \hat{X}_i)u_i(t, X)D) dt. \qquad (\dagger)$$

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• Let  $X^{(i)}$  denote the random vector  $(X_1, \ldots, X_{i-1}, 0, X_{i+1}, \ldots, X_n)$  and let  $D_i := h(X^{(i)})$ .

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- ► Then note that u<sub>i</sub>(t, X<sup>(i)</sup>)D<sub>i</sub> is a function of the random variables (X<sub>j</sub>)<sub>j≠i</sub> only.

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- Therefore since  $\hat{X}_i = \mathbb{E}(X_i \mid (X_j)_{j \neq i})$ ,

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Thus,

$$egin{aligned} \mathbb{E}((X_i-\hat{X}_i)u_i(t,X)D)\ &=\mathbb{E}((X_i-\hat{X}_i)u_i(t,X)D)-\mathbb{E}((X_i-\hat{X}_i)u_i(t,X^{(i)})D_i)\,. \end{aligned}$$

- ► Let  $X^{(i)}$  denote the random vector  $(X_1, \ldots, X_{i-1}, 0, X_{i+1}, \ldots, X_n)$  and let  $D_i := h(X^{(i)})$ .
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If the smoothness term is small, then u<sub>i</sub>(t, X) ≈ u<sub>i</sub>(t, X<sup>(i)</sup>) and D ≈ D<sub>i</sub>. Together with the identity (†), this shows that f(X) ≈ f(X̂) with high probability.

 Large deviations for sparse random graphs cannot be tackled by techniques based on graph limit theory and Szemerédi's regularity lemma.

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- Large deviations for sparse random graphs cannot be tackled by techniques based on graph limit theory and Szemerédi's regularity lemma.
- The recently developed theory of large deviations for nonlinear functions of Bernoulli random variables has seen some successful applications in this class of problems, leading to the solutions of some longstanding questions.

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- Large deviations for sparse random graphs cannot be tackled by techniques based on graph limit theory and Szemerédi's regularity lemma.
- The recently developed theory of large deviations for nonlinear functions of Bernoulli random variables has seen some successful applications in this class of problems, leading to the solutions of some longstanding questions.
- The degree of sparsity that is allowed by these theorems is less than optimal. New breakthroughs are required.

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