### Local Privacy and Statistical Minimax Rates

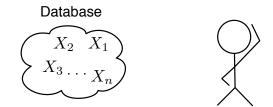
John C. Duchi, Michael I. Jordan, Martin J. Wainwright

University of California, Berkeley

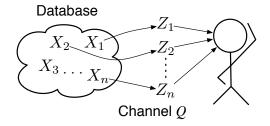
December 2013

Goals for this talk

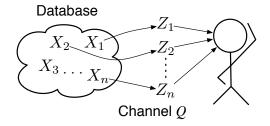
### Bring together some classical concepts of decision theory and newer concepts of privacy



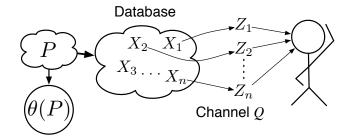
• Have data  $X_1, \ldots, X_n$ 



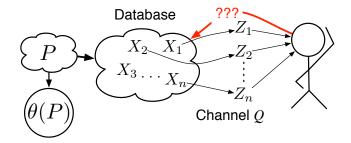
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- Private views  $Z_1, \ldots, Z_n$  constructed from  $X_i$



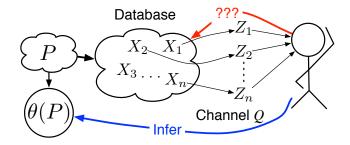
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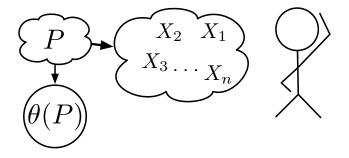


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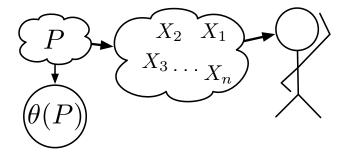


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- Goal: infer population parameter  $\theta(P)$  based on  $Z_1, \ldots, Z_n$

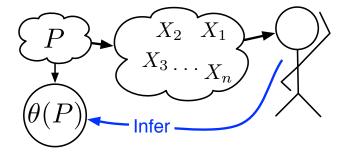
# Primer on minimax rates of convergence and statistical inference



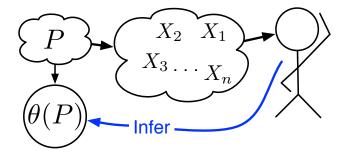
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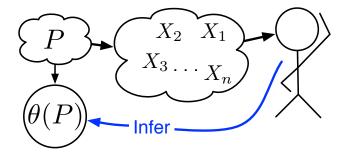


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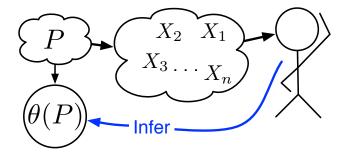
Why? Care about making future predictions



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Why? Care about making future predictions

What is likelihood new resident of San Francisco needs food stamps



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Why? Care about making future predictions

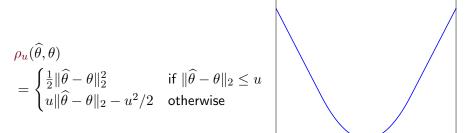
- What is likelihood new resident of San Francisco needs food stamps
- Biological prediction, web advertising, search, …

#### Central object of study: Minimax risk

- Parameter  $\theta(P)$  of distribution P
  - E.g. mean:  $\theta(P) = \mathbb{E}_P[X]$

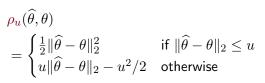
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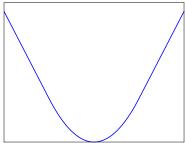
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  - ▶ E.g.  $\rho(\hat{\theta}, \theta) = \|\hat{\theta} \theta\|_2^2$  or more esoteric/robust



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- Family of distributions *P* that we study
  - E.g. P such that  $\mathbb{E}_P[X^2] \leq 1$

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$$\inf_{\widehat{\theta}} \sup_{P \in \mathcal{P}} \mathbb{E}_P \left[ \rho(\widehat{\theta}(X_1, \dots, X_n), \theta(P)) \right]$$

- Worst case over distributions  ${\cal P}$
- Best case over all estimators  $\widehat{\theta} : \mathcal{Z}^n \to \Theta$

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To study: rate of  $\mathfrak{M}_n(\theta(\mathcal{P}), \rho) \to 0$  as n grows

This talk:

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- Upper bounds will be ad-hoc
- Lower bounds will be information theoretic [Hasminskii 78, Birge 83, Ibragimov and Hasminskii 81, Yang and Barron 99, Yu97]

#### This talk:

- Upper bounds will be ad-hoc
- Lower bounds will be information theoretic [Hasminskii 78, Birge 83, Ibragimov and Hasminskii 81, Yang and Barron 99, Yu97]
- ▶ **NB:** Many known information-theoretic upper bounds [Barron, Birge, Kivinen, ...]

Step 1: Reduce from estimation to testing

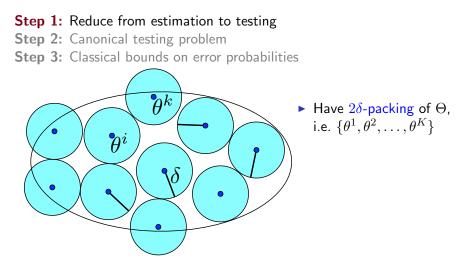
- Step 1: Reduce from estimation to testing
- Step 2: Canonical testing problem

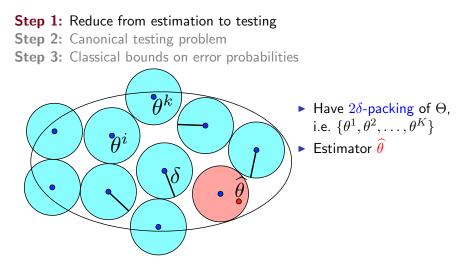
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#### Step 1: Reduce from estimation to testing

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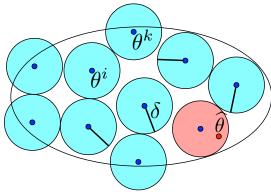




# **Step 1:** Reduce from estimation to testing Step 2: Canonical testing problem Step 3: Classical bounds on error probabilities $\mathbf{A}^k$ • Have $2\delta$ -packing of $\Theta$ , i.e. $\{\theta^1, \theta^2, \dots, \theta^K\}$ $\theta^i$ • Estimator $\hat{\theta}$ At most one index close to $\hat{\theta}$

#### Step 1: Reduce from estimation to testing

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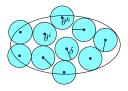


- ► Have 2δ-packing of Θ, i.e. {θ<sup>1</sup>, θ<sup>2</sup>,...,θ<sup>K</sup>}
- Estimator  $\hat{\theta}$
- At most *one* index close to θ
- Can *test* index  $i \in [K]$

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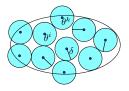
Step 2: Canonical testing problem

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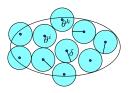
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- Lower bound minimax error:

$$\sup_{P \in \mathcal{P}} \mathbb{E}_{P} \left[ \rho(\widehat{\theta}, \theta(P)) \right] \geq \frac{1}{K} \sum_{v=1}^{K} \mathbb{E}_{v} \left[ \rho(\widehat{\theta}, \theta_{v}) \right]$$
$$\geq \frac{1}{K} \sum_{v=1}^{K} \rho(\delta) P_{v}(\rho(\widehat{\theta}, \theta_{v}) \geq 2\delta)$$

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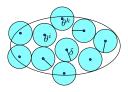
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Final canonical testing problem:

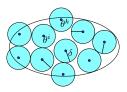
$$\mathfrak{M}_n(\theta(\mathcal{P}), \rho) \ge \rho(\delta) \min_{\widehat{v}} \mathbb{P}(\widehat{v}(X_1, \dots, X_n) \neq V).$$

Step 1: Reduce from estimation to testing

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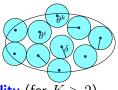
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$$P_0(\hat{v} \neq 0) + P_1(\hat{v} \neq 1)$$
  
 
$$\geq 1 - \|P_0 - P_1\|_{\mathrm{TV}}$$

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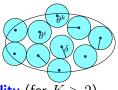
**Fano's inequality** (for  $\bar{K} > 2$ )

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$$\geq 1 - \frac{I(X_1, \dots, X_n; V) + \log 2}{\log K}$$

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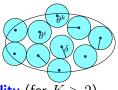
#### Le Cam's method

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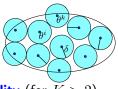
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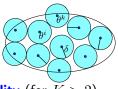
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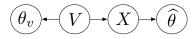


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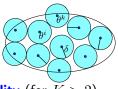
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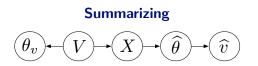
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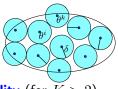
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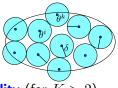
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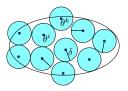
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Summarizing  

$$(\theta_v) \leftarrow V \rightarrow X \rightarrow \widehat{\theta} \rightarrow \widehat{v}$$
  
 $\mathfrak{M}_n(\theta(\mathcal{P}), \rho) \ge \rho(\delta) \min \mathbb{P}(\widehat{v}(X_1, \dots, X_n) \neq V)$ 

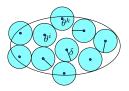
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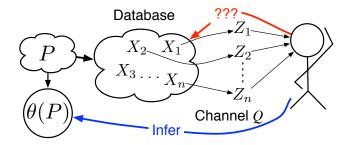
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Key idea: Control information-theoretic divergences

$$||P_0 - P_1||_{\text{TV}}$$
 or  $I(X_1, \dots, X_n; V)$ 

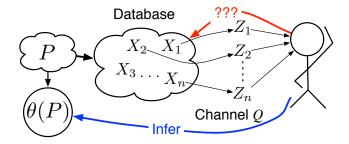
to attain minimax rate

#### Inference under privacy constraints



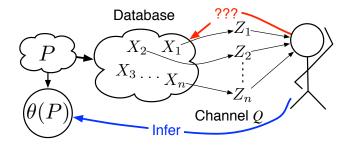
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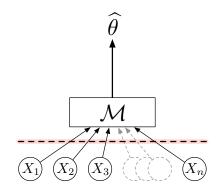


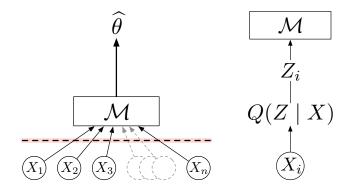
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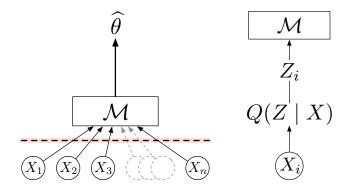


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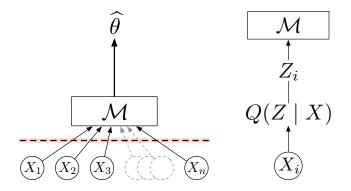




**Local Privacy:** Don't trust collector of data (Evfimievski et al. 2003, Warner 1965)



• Individuals  $i \in \{1, \ldots, n\}$  have personal data  $X_i \sim P_{\theta}$ 

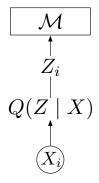


- ► Individuals  $i \in \{1, ..., n\}$  have personal data  $X_i \sim P_\theta$
- Estimator  $Z_1^n \mapsto \widehat{\theta}(Z_{1:n})$

**Definition:** The channel Q is  $\alpha$ -differentially private if

$$\max_{z,x,x'} \frac{Q(Z=z \mid x)}{Q(Z=z \mid x')} \le e^{\alpha}.$$

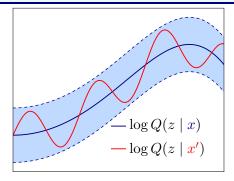
[Dwork, McSherry, Nissim, Smith 2006]

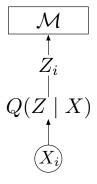


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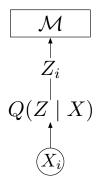
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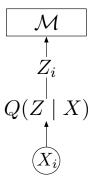
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$$\mathsf{FNR} + \mathsf{FPR} \geq \frac{2}{1 + e^\alpha}$$

# $\begin{array}{c|c} \mathcal{M} \\ \uparrow \\ Z_i \\ Q(Z \mid X) \\ \uparrow \\ X_i \end{array}$

#### [Wasserman and Zhou 2011]

#### Central object of study: Minimax risk

- Parameter  $\theta(P)$  of distribution P
- Loss  $\rho$  that measures error in estimate of  $\hat{\theta}$  for  $\theta$ :  $\rho(\hat{\theta}, \theta)$

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$$\rho(\widehat{\theta}, \theta) = \|\widehat{\theta} - \theta\|_2^2$$

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#### Many related results

 Non-population lower bounds [Hardt and Talwar 10, Nikkolov, Talwar, Zhang 13; Hall, Rinaldo, Wasserman 11, Chaudhuri, Monteleoni, Sarwate 12]

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# Examples

- Mean estimation
- Fixed-design regression
- Convex risk minimization (i.e. online learning)
- Multinomial (probability) estimation
- Nonparametric density estimation

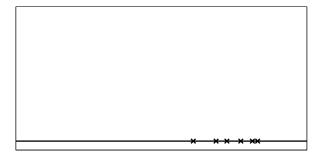
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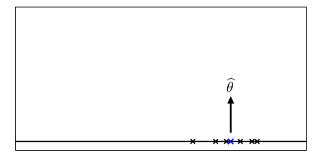
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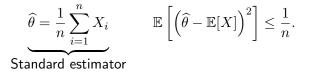
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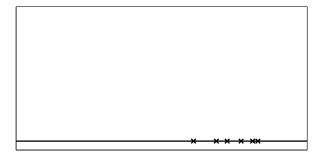
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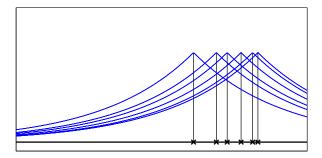
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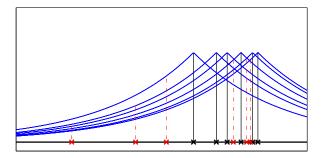
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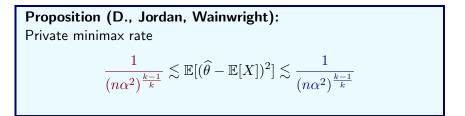


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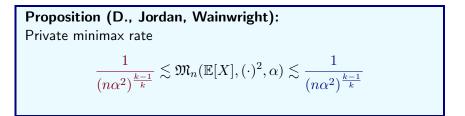
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Proposition (D., Jordan, Wainwright): Non-private minimax rate $\frac{1}{n}\lesssim \mathbb{E}[(\widehat{\theta}-\mathbb{E}[X])^2]\lesssim \frac{1}{n}$ 

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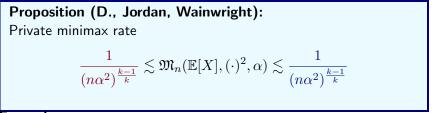


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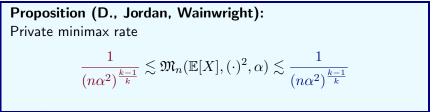


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- For  $k \to \infty$  (bounded random variables) parametric decrease

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	0701

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Proposition: Non-private minimax rate

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Take away: Sample size reduction

$$n\mapsto rac{nlpha^2}{d}$$

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### Example 2: multinomial estimation

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## Main consequences

**Goal:** Understand tradeoff between differential privacy bound  $\alpha$  and sample size n

**"Theorem 1"** Effective sample size for *essentially any*<sup>1</sup> problem is made worse by at least

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**"Theorem 2"** Effective sample size for *d*-dimensional problems scales as

$$n \mapsto \frac{n\alpha^2}{d}$$

## General theory

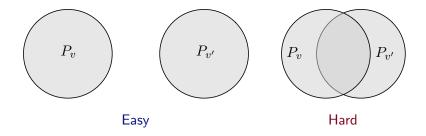
#### Showing minimax bounds:

- Have possible "true" parameters  $\{\theta_v\}$  we want to find
- Distribution  $P_v$  associated with each parameter
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### Differential privacy and probability distributions

Samples:  $Z_i$  are drawn  $X_i \to Q \to Z_i$  from marginal  $M_v(Z) := \int Q(Z \mid X = x) dP_v(x)$ 

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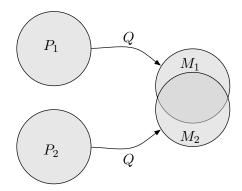
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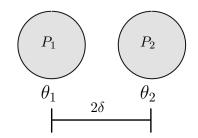


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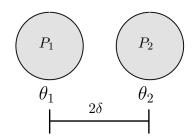


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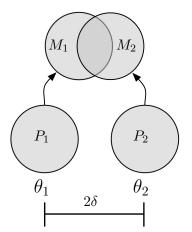
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$$\mathfrak{M}_{n}\left(\Theta, (\cdot)^{2}, \alpha\right)$$

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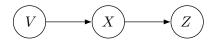


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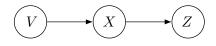
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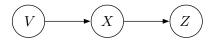
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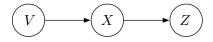


**Goal:** Find V based on  $Z_1, \ldots, Z_n$ **Difficulty of problem:** Saw earlier *mutual information* 

$$I(X_1, \ldots, X_n; V) \mapsto I(Z_1, \ldots, Z_n; V)$$

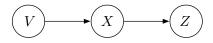


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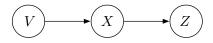


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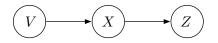


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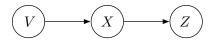
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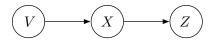
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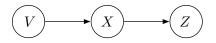
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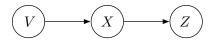
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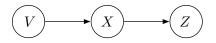
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**Pre/post-print online:** "Local privacy and statistical minimax rates." D., Jordan, & Wainwright (2013). *arXiv:1302.3203 [stat.TH]*